

# Regularly Varying Random Fields

## BIRS-CMO Workshop

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# Multivariate Regular Variation

## Definition.

A random vector  $\mathbf{X}$  is said to be regularly varying with index  $\alpha \in (0, \infty)$  if there exists a regularly function  $V(x)$  with index  $-\alpha$  (i.e.,  $\lim_{x \rightarrow \infty} V(ux)/V(x) = u^{-\alpha}$  for  $u > 0$ ), and a nonzero Radon measure  $\mu$  on  $\bar{\mathbb{R}}^d \setminus \{\mathbf{0}\} = [-\infty, \infty]^d \setminus \{\mathbf{0}\}$  such that

$$\frac{\mathbb{P}(x^{-1}\mathbf{X} \in \cdot)}{V(x)} \xrightarrow{v} \mu(\cdot)$$

as  $x \rightarrow \infty$ .

# Tail Dependence of Stationary Stochastic Processes

- Leadbetter (1983): extremal index  $\theta$

Assuming the regular variation...

- Basrak and Segers (2009): tail process

$$\mathcal{L}(x^{-1}\mathbf{X}_s, \dots, x^{-1}\mathbf{X}_t \mid \|\mathbf{X}_0\| > x) \xrightarrow{x \rightarrow \infty} \mathcal{L}(\mathbf{Y}_s, \dots, \mathbf{Y}_t \mid \|\mathbf{X}_0\| > x)$$

# Questions

- Can we go from  $\mathbf{X}_t, t \in \mathbb{Z}$  to  $\mathbf{X}(\mathbf{t}), \mathbf{t} \in \mathbb{Z}^k$ ?
- What is an appropriate way to extend?

# Outline

- 1 Introduction
- 2 The Spatial Extremal Index
- 3 The Tail Field
- 4 Relationship Between  $\mathbf{Y}(\mathbf{t})$  And  $\theta$
- 5 Application: Brown-Resnick Random Fields (BRRFs)

# Notations

- $M_X(A) = \max_{\mathbf{t} \in A} \|\mathbf{X}(\mathbf{t})\|$ , for an index set  $A \subset \mathbb{Z}^k$
- $\mathcal{R}_{\mathbf{n}} = \{\mathbf{i} \in \mathbb{Z}^k : -(n_\ell - 1) \leq i_\ell \leq (n_\ell - 1), \ell = 1, \dots, k\}$
- $\mathcal{R}_{\mathbf{n}}^+ = \{\mathbf{i} \in \mathbb{Z}^k : 0 \leq i_\ell \leq (n_\ell - 1), \ell = 1, \dots, k\}$

# The Spatial Extremal Index

## Definition.

Let  $(\mathbf{X}(\mathbf{t}) : \mathbf{t} \in \mathbb{Z}^k)$  be a stationary random field, and  $(\mathcal{R}_n)$  be a sequence of nondecreasing hypercubes. Suppose for each fixed  $\tau > 0$  and any sequence  $(u_n(\tau))$  satisfying

$$|\mathcal{R}_n| \mathbb{P}(\|\mathbf{X}(\mathbf{0})\| > u_n(\tau)) \rightarrow \tau$$

as  $n \rightarrow \infty$ , it holds that

$$\mathbb{P}(M_X(\mathcal{R}_n) \leq u_n(\tau)) \rightarrow e^{-\theta\tau}.$$

Then we say that the extremal index of the random field is  $\theta$ .

# Computing $\theta$ : Method 1

- Stationary stochastic processes

$$\theta = \lim_{n \rightarrow \infty} \mathbb{P} \left( \max_{i=1, \dots, r_n-1} \|\mathbf{X}_i\| \leq u_n \mid \|\mathbf{X}_0\| > u_n \right)$$

(O'Brien (1987))



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(O'Brien (1987))

- Stationary random fields

?

For stationary stochastic processes:

$$\begin{aligned} & \mathbb{P} \left( \max_{i=1, \dots, r_n-1} \|\mathbf{X}_i\| \leq u_n \mid \|\mathbf{X}_0\| > u_n \right) \\ &= \mathbb{P} \left( \max_{i=0, \dots, r_n-2} \|\mathbf{X}_i\| \leq u_n \mid \|\mathbf{X}_{r_n-1}\| > u_n \right) \end{aligned}$$

For random fields:

Which corner should we condition on?

## Example

- $Z(\mathbf{t}), \mathbf{t} \in \mathbb{Z}^2$ : iid standard Pareto random variables
- $X(\mathbf{t}) = \max\{Z(\mathbf{t} - \mathbf{1}), Z(\mathbf{t})\}$

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## Computing $\theta$ : Method 2

- Stationary stochastic processes, assuming vanish condition

$$\theta = \lim_{n \rightarrow \infty} \frac{\mathbb{P}(M_X(\mathcal{R}_{r_n}^+) > u_n)}{r_n \mathbb{P}(\|\mathbf{X}_0\| > u_n)} \quad (\text{Basrak and Segers (2009)})$$

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- Stationary random fields, assuming  $\Delta(u_n)$ -condition

$$\begin{aligned} \theta &= \lim_{\mathbf{n} \rightarrow \infty} \left( \mathbb{E} \left[ \sum_{\mathbf{t} \in \mathcal{R}_{r_n}^+} \mathbb{1}(\|\mathbf{X}(\mathbf{t})\| > u_n) \mid M_X(\mathcal{R}_{r_n}^+) > u_n \right] \right)^{-1} \\ &= \lim_{\mathbf{n} \rightarrow \infty} \frac{\mathbb{P}(M_X(\mathcal{R}_{r_n}^+) > u_n)}{\left( \prod_{\ell=1}^k r_{n_\ell} \right) \mathbb{P}(\|\mathbf{X}(\mathbf{0})\| > u_n)} \quad (\text{Pereira et al (2017)}) \end{aligned}$$

# Condition: Vanish Condition

- $\lim_{\mathbf{m} \rightarrow \infty} \limsup_{\mathbf{n} \rightarrow \infty} \mathbb{P}(M_X(\mathcal{R}_{r_n} \setminus \mathcal{R}_m) > u_n \mid \|\mathbf{X}(\mathbf{0})\| > u_n) = 0$
- $\mathbf{Y}(\mathbf{t}) \xrightarrow{\text{a.s.}} \mathbf{0}$  as  $\mathbf{t} \rightarrow \infty$ ,  $\theta > 0$

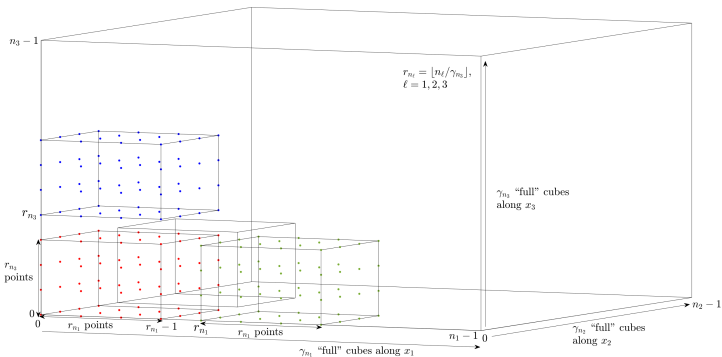
# Condition: $\Delta(u_n)$ -condition

- The coordinatewise mixing condition (Choi (2002))



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# Existence Of The Tail Field

## Theorem 1.

Let  $(\mathbf{X}(\mathbf{t}) : \mathbf{t} \in \mathbb{Z}^k)$  be a stationary random field, and  $\alpha \in (0, \infty)$ . Then it is jointly regularly varying with index  $\alpha$ , if and only if there exists a random field  $(\mathbf{Y}(\mathbf{t}) : \mathbf{t} \in \mathbb{Z}^k)$  such that

$$\mathcal{L}(x^{-1}\mathbf{X}(\mathbf{t}) : \mathbf{t} \in \mathbb{Z}^k \mid \|\mathbf{X}(\mathbf{0})\| > x) \rightarrow \mathcal{L}(\mathbf{Y}(\mathbf{t}) : \mathbf{t} \in \mathbb{Z}^k)$$

as  $x \rightarrow \infty$ , and  $\mathbb{P}(\|\mathbf{Y}(\mathbf{0})\| > y) = y^{-\alpha}$  for  $y \geq 1$ .

# Properties Of The Tail Field

① Let  $\Theta(\mathbf{t}) = \mathbf{Y}(\mathbf{t})/\|\mathbf{Y}(\mathbf{0})\|$ , then  $\|\mathbf{Y}(\mathbf{0})\| \perp (\Theta(\mathbf{t}) : \mathbf{t} \in \mathbb{Z}^k)$ .

② For any bounded measurable  $g : (\overline{\mathbb{R}}^d)^{\mathbb{Z}^k} \rightarrow \mathbb{R}$ , and  $\mathbf{s} \in \mathbb{Z}^k$ :

- $\mathbb{E}[g(\mathbf{Y}(\cdot - \mathbf{s}))\mathbb{1}(\mathbf{Y}(-\mathbf{s}) \neq \mathbf{0})]$   
 $= \int_0^\infty \mathbb{E}[g(r\Theta(\cdot))\mathbb{1}(r\|\Theta(\mathbf{s})\| > 1)] d(-r^{-\alpha})$
- $\mathbb{E}[g(\Theta(\cdot - \mathbf{s}))\mathbb{1}(\Theta(-\mathbf{s}) \neq \mathbf{0})]$   
 $= \mathbb{E}\left[g\left(\frac{\Theta(\cdot)}{\|\Theta(\mathbf{s})\|}\right)\|\Theta(\mathbf{s})\|^\alpha\right]$

## The Sufficient Condition for The Joint Regular Variation

- Stochastic Processes  
the weak convergence on the set of nonnegative times  
 $\Rightarrow$  joint regular variation of the original process  
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  - Random Fields  
the weak convergence on the set in first 'quadrant'  
 $\stackrel{?}{\Rightarrow}$  joint regular variation of the original random field
- No!
- The weak convergence of the **entire** random field is needed.

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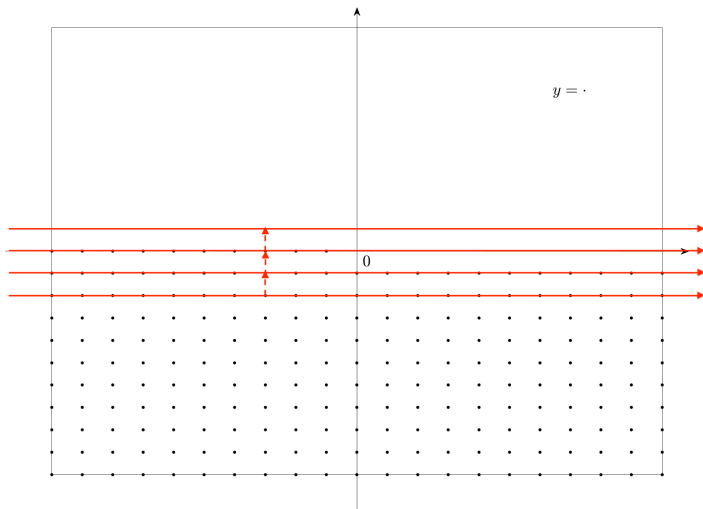


# Computing $\theta$ : The Tail Field Expression

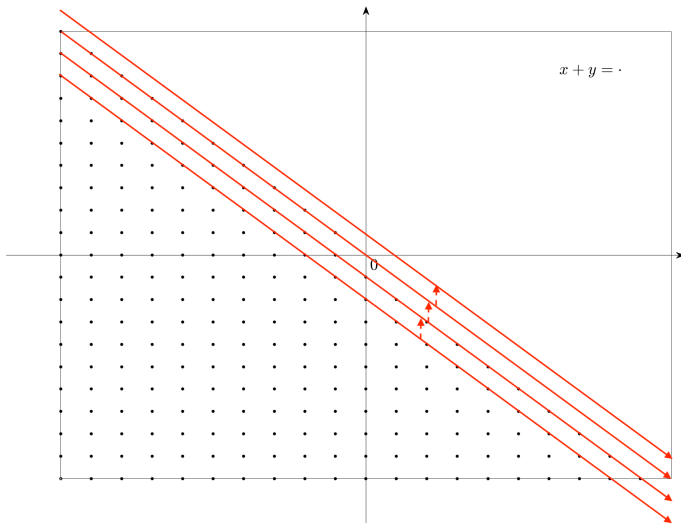
$(\mathbf{X}(\mathbf{t}) : \mathbf{t} \in \mathbb{Z}^k)$ : stationary, satisfying  $\Delta(u_n)$ -condition, and vanish condition.

- $\theta = \mathbb{P}(\max_{\mathbf{t} \prec \mathbf{0}} \|\mathbf{Y}(\mathbf{t})\| \leq 1)$

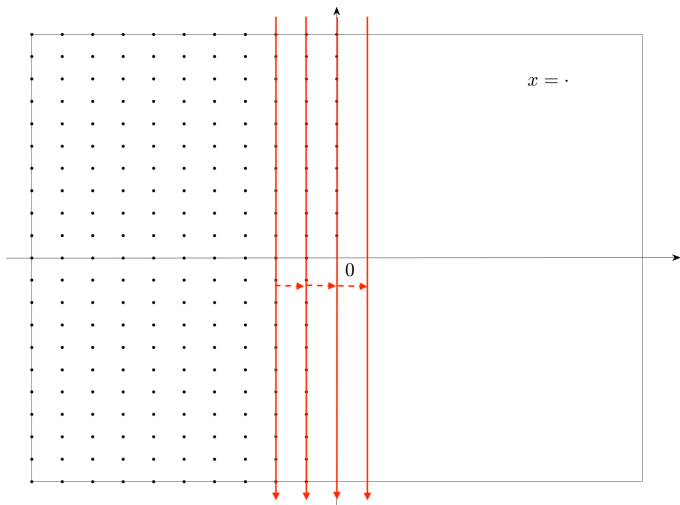
# Ordering Examples in 2D



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- $\theta = \mathbb{P}(\max_{\mathbf{t} \prec \mathbf{0}} \|\mathbf{Y}(\mathbf{t})\| \leq 1)$
- Ordering does not matter!

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# Constructing BRRFs

- $(W(\mathbf{t}) : \mathbf{t} \in \mathbb{Z}^k)$ : zero-mean Gaussian random field  
 $\mathcal{L}(W(\mathbf{t} + \mathbf{s}) - W(\mathbf{s}))$  does not depend on  $\mathbf{s}$
- $(W_i(\mathbf{t}) : \mathbf{t} \in \mathbb{Z}^k), i \in \mathbb{N}$ : iid copies of  $(W(\mathbf{t}) : \mathbf{t} \in \mathbb{Z}^k)$
- $\sum_{i=1}^{\infty} \delta_{U_i}$ : a Poisson process on  $\mathbb{R}$  with intensity  $du/u^2$

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- $\sum_{i=1}^{\infty} \delta_{U_i}$ : a Poisson process on  $\mathbb{R}$  with intensity  $du/u^2$
- $X(\mathbf{t}) = \max_{i=1,2,\dots} U_i \exp\{W_i(\mathbf{t}) - \sigma^2(\mathbf{t})/2\}$



# Its Tail Field

- The Joint Distribution

$$\begin{aligned} & \mathbb{P}(Y(\mathbf{t}_1) < y_1, \dots, Y(\mathbf{t}_n) < y_n) \\ &= \mathbb{E} \left[ \max_{i=1, \dots, n} \left( \frac{1}{y_i} \exp \left\{ W(\mathbf{t}_i) - \frac{\sigma^2(\mathbf{t}_i)}{2} \right\}, \exp \left\{ W(\mathbf{0}) - \frac{\sigma^2(\mathbf{0})}{2} \right\} \right) \right] \\ &= \mathbb{E} \left[ \max_{i=1, \dots, n} \frac{1}{y_i} \exp \left\{ W(\mathbf{t}_i) - \frac{\sigma^2(\mathbf{t}_i)}{2} \right\} \right] \end{aligned}$$

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- The Marginal Distribution

$$\mathbb{P}(Y(\mathbf{t}) < y) = \Phi \left( \frac{\ln y + \gamma(\mathbf{t})}{\sqrt{2\gamma(\mathbf{t})}} \right) - \frac{1}{y} \Phi \left( \frac{\ln y - \gamma(\mathbf{t})}{\sqrt{2\gamma(\mathbf{t})}} \right)$$

# The Spatial Extremal Index

Recall:  $\theta = \mathbb{P}(\max_{\mathbf{t} \prec \mathbf{0}} Y(\mathbf{t}) \leq 1)$

## Corollary 2.

*Let  $(X(\mathbf{t}) : \mathbf{t} \in \mathbb{Z}^k)$  be a BRRF with standard Fréchet margin, satisfying both  $\Delta(u_n)$ -condition and vanish condition. Then*

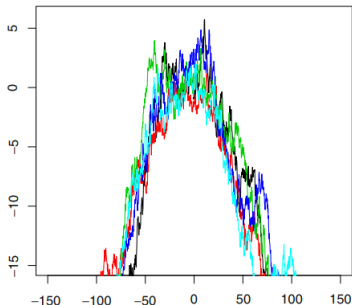
$$\theta = \mathbb{E} \left[ \max_{\mathbf{t} \preceq \mathbf{0}} \exp \{ W(\mathbf{t}) - \sigma^2(\mathbf{t})/2 \} \right] - \mathbb{E} \left[ \max_{\mathbf{t} \prec \mathbf{0}} \exp \{ W(\mathbf{t}) - \sigma^2(\mathbf{t})/2 \} \right].$$

In comparison with

$$\theta = \lim_{n \rightarrow \infty} \left( \mathbb{E} \left[ \sum_{\mathbf{t} \in \mathcal{R}_{r_n}^+} \mathbb{1}(X(\mathbf{t}) > u_n) \mid M_X(\mathcal{R}_{r_n}^+) > u_n \right] \right)^{-1}$$

# Simulating BRRFs Is Hard

- Approximation of the Brown-Resnick process based on the definition may result in non-stationarity



Oesting et al (2012). Simulation of Brown-Resnick processes.

# Example: Brownian Motions

- $(W_1(t) : t \in \mathbb{Z}), (W_2(t) : t \in \mathbb{Z})$ : two independent two-sided standard Brownian motions
- $W(t_1, t_2) = W_1(t_1) + W_2(t_2)$

Order	$y = \cdot$	$x + 3y = \cdot$	$x + 2y = \cdot$	$x + y = \cdot$
$\hat{\theta}$	0.0780	0.0786	0.0786	0.0784
$sd(\hat{\theta})$	0.0008	0.0008	0.0008	0.0008
Order	$2x + y = \cdot$	$3x + y = \cdot$	$x = \cdot$	
$\hat{\theta}$	0.0786	0.0787	0.0781	
$sd(\hat{\theta})$	0.0008	0.0008	0.0008	

Table: Simulated  $\theta$  with different ordering

Thank you!