

Change point and trend analyses of annual expectile curves of tropical storms

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Joint work with

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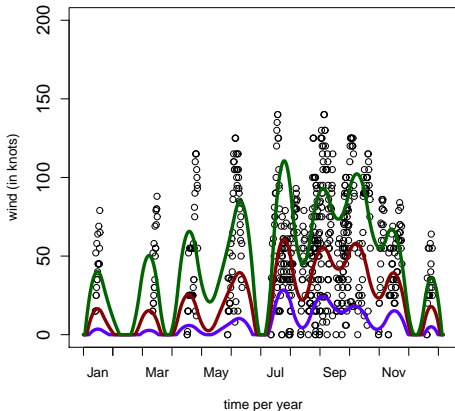
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Outline of the talk:

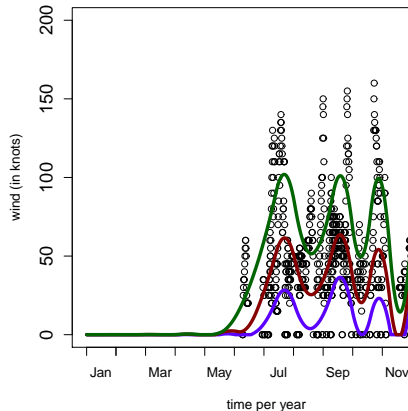
- ① Expectile curves of tropical storm strength
- ② A functional regression model
- ③ Change point and trend tests
- ④ Application to hurricane and typhoon data

2005 expectile curves

Typhoons in 2005



Hurricanes in 2005



Expectile curves for $\tau = 0.1, 0.5$ and 0.9 .

Expectiles, background

Y is a square integrable random variable

$$E_-(e) = E [(X - e)^2 \mathbf{I}\{X \leq e\}] ;$$

$$E_+(e) = E [(X - e)^2 \mathbf{I}\{X > e\}] .$$

$$E_\tau(e) = (1 - \tau)E_-(e) + \tau E_+(e),$$

The τ th expectile e_τ minimizes $E_\tau(e)$.

(τ close to 1, $E_+(e)$ must be small, e_τ must be large.)

Estimation through empirical expectations.

Idea can be extended to the scatter plot of points (t_i, x_i) , $t_i \in I$.

Our application: $I = \text{year}$, x_i wind speed at time t_i .

Using spline smoothing, one obtains expectile curves.

change point test

We observe curves X_1, X_2, \dots, X_N , $\mu_i(t) = EX_i(t)$.

$$H_0 : \mu_1 = \mu_2 = \dots = \mu_N.$$

Test based on projections of

$$P_k(t, \tau) = \frac{k(N-k)}{N} \{ \hat{\mu}_k(t, \tau) - \tilde{\mu}_k(t, \tau) \},$$

Scores:

$$\hat{\xi}_{j,n} = \int \{ X_n(t) - \bar{X}_N(t) \} \hat{v}_j(t) dt.$$

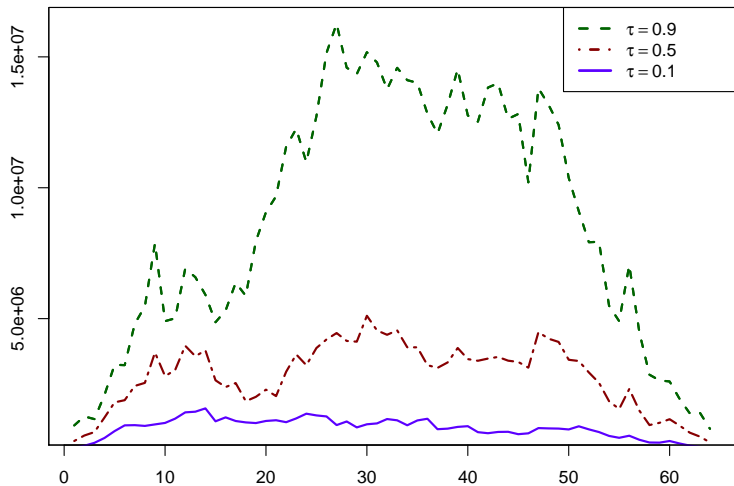
Test statistics:

$$\hat{S}_d = \frac{1}{N^2} \sum_{j=1}^d \frac{1}{\hat{\lambda}_j} \sum_{k=1}^N \left(\sum_{1 \leq i \leq k} \hat{\xi}_{j,i} - \frac{k}{N} \sum_{1 \leq i \leq k} \hat{\xi}_{j,i} \right).$$

Limit distribution of Kiefer's type

(sum of integrals of squared Brownian bridges)

$P_n(t; \tau)$ functions, typhoons



$$X_n(t) = \alpha(t) + \beta(t)n + \varepsilon_n(t).$$

$$H_0 : \beta = 0, \quad \text{vs.} \quad H_A : \beta \neq 0.$$

LSE:

$$\hat{\beta}(t) = \frac{6}{N(N+1)(N-1)} \sum_{k=1}^N (2k - N - 1)X_k(t)$$

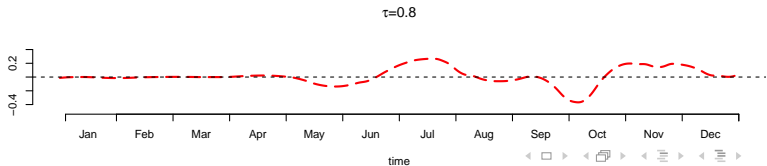
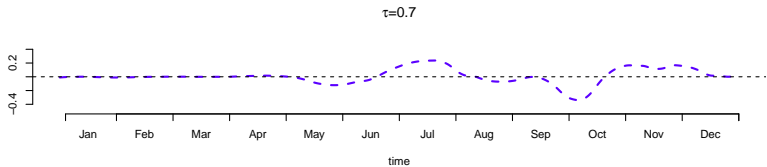
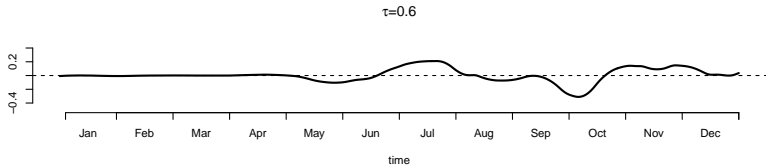
If the function $\hat{\beta}$ is large, we reject H_0 .

λ_j eigenfunctions of the error curves ε_i

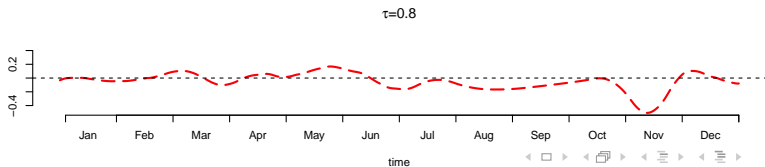
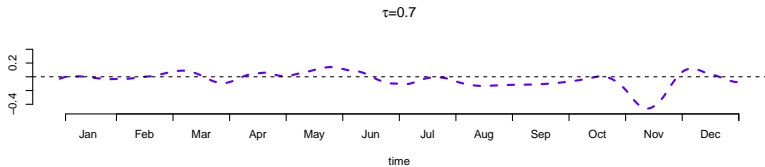
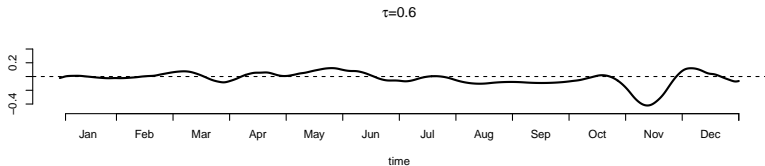
$\hat{\lambda}_j$ sample eigenvalues of the residual curves $\hat{\varepsilon}_i$

\hat{v}_j corresponding eigenfunctions (sample FPC's)

Hurricanes coefficient functions

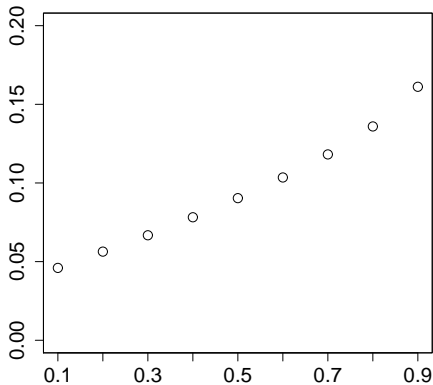


Typhoons coefficient functions

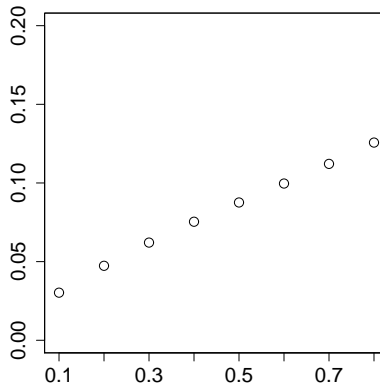


Norms of the slope functions $\hat{\beta}$

Typhoons



Hurricanes



Trend tests

Monte Carlo test:

Under H_0 ,

$$\hat{\Lambda}_N = \frac{N^3}{12} \int_0^1 (\hat{\beta}(t))^2 dt \xrightarrow{\mathcal{L}} \Lambda_\infty \stackrel{\text{def}}{=} \sum_{j=1}^{\infty} \lambda_j Z_j^2,$$

(Z_j independent standard normal)

Chi-square test:

Under H_0 ,

$$\hat{T}_N = \frac{N^3}{12} \sum_{j=1}^q \hat{\lambda}_j^{-1} \langle \hat{\beta}, \hat{v}_j \rangle^2 \xrightarrow{\mathcal{L}} \chi_q^2.$$

Both tests are shown to be consistent.

Application to typhoon and hurricane data

Change point test rejects H_0 for all values of $\tau \in .1, .2, \dots, .9$

Monte Carlo test P-values

τ	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9
typhoons	0.365	0.537	0.545	0.495	0.438	0.381	0.329	0.300	0.300
hurricanes	0.439	0.239	0.133	0.081	0.062	0.047	0.038	0.030	0.030

Chi-square test P-values

τ	0.1	0.2	0.3	0.4	0.5	0.6	0.7	0.8	0.9
q	10	11	12	12	12	12	12	12	12
typhoons	0.534	0.705	0.722	0.688	0.587	0.466	0.382	0.300	0.300
q	5	5	5	6	6	6	7	7	7
hurricanes	0.069	0.024	0.015	0.006	0.003	0.003	0.004	0.004	0.004

Main conclusions

- Simulations show that the Monte Carlo test is more accurate if DGP's resemble actual data.
- The annual pattern of wind speeds of both hurricanes and typhoons has been changing at all wind speed levels over the last 60 years.
- There is a significant trend in the shape of this pattern for upper wind speed levels of hurricanes.