

# **Smooth modeling of covariate effects in bisulfite sequencing-derived measures of DNA methylation**

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# Epigenetics and DNA Methylation



- ▶ change gene expression without changing DNA sequence
- ▶ can be altered by age, diet, stress and environmental exposures



- ▶ **interest:** identify **genomic regions** where DNA methylation patterns demonstrate alterations associated with disease phenotypes (DMRs).

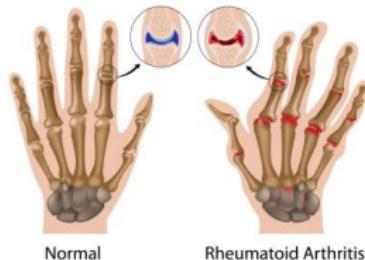
# Motivating Dataset

Methylation profiles of Rheumatoid Arthritis (RA) patients and controls  
(provided courtesy of Dr. Marie Hudson, McGill)



- **Samples:** cell-separated blood samples of 22 RA patients and 21 healthy individuals from either T cells or monocytes.

	MONO	TCELL
RA	10	12
Control	8	13



- **Methylation:** targeted custom captured bisulfite sequencing †
  - Prior selection of predefined genomic regions of interest
  - ~ 400, 000 regions in the genome
- This presentation focuses on **one targeted region**
  - **Chr4: 102,711,629 – 102,712,832** (near gene *BANK1*)
  - consists of 123 CpG sites

† Allum et al (2015) *Nature*.

# Quantification – Bisulfite sequencing



Position	Unmeth counts	Meth counts	Total counts	Sample ID	Sample-level covariates
102711629	2	2	4	1	...
102711630	15	0	15	1	
102711649	15	0	15	1	
102711650	8	0	8	1	
102711850	15	0	15	2	
102711851	4	9	13	2	

**Sample-level covariates:** disease status, cell type composition, age, smoking status...

# Challenges



## ► Read-depth variability

- The total number of reads varies at different CpG sites. Modeling the proportion treats noisy measurements the same way as accurate ones ( $\frac{5}{10} = \frac{50}{100}$ )
- missing values occur frequently

## ► Variability in cell-type mixture proportions

→ adjusting for multiple covariates

- Methylation levels vary substantially across different cell types, which can confound the association of interest.

## ► Experimental errors

- Sequencing errors: more mis-alignment of unmethylated reads after bisulfite sequencing
- Bisulfite conversion errors: incomplete C-T conversion; or over-treatment with bisulfite leading to conversion of methylated C to T

# Existing methods & Motivations



Challenges	BSmooth (Hansen et al., 2012)	BiSeq (Hebestreit et al., 2013)	SMSC (Lakhal-Chaieb et.al., 2017)
Variable read-depth	✓		✓
Experimental error			✓
Mixture of cell types		✓	
Multiple covariates		✓	

- ▶ Most of the existing methods are of **two-stage** nature
  - (1) smooth the methylated proportions **for each sample**, and
  - (2) fit model (t-test or beta regression) to the **smoothed methylation data**.
- ▶ **Motivation:** extend the work in Lakhal-Chaieb et.al., 2017 to enable an integrated analysis of **multiple samples** that allows for
  - **cell type mixtures**, and
  - **multiple covariates** in the model.

# Notation & Model

Let  $(i, j, k)$  index sample, CpG sites and reads respectively.

$$i = 1, 2, \dots, I; \quad j = 1, 2, \dots, n_i; \quad k = 1, 2, \dots, X_{ij}, \quad N = \sum_{i=1}^I n_i.$$

Example: CpG site  $j$  for Individual  $i$

	$t_{ij}$	$Y_{ij}$	$X_{ij}$	$Z_{1i}, Z_{2i}, \dots, Z_{Pi}$ Sample-level covariates			
(1, 1)	114354051	2	4	1 ...	M	$Y_{ij1} = 1$	$S_{ij1}$
(1, 2)	114354052	0	15	1	U	$Y_{ij2} = 0$	$S_{ij2}$
...	114354053	0	15	1	U	$Y_{ij3} = 0$	$S_{ij3}$
(1, n1)	114354054	0	14	1	M	$Y_{ij4} = 1$	$S_{ij4}$
(2, 1)	114354056	0	15	2			
(2, 2)	114354057	9	13	2			
						$\sum Y_{ijk} = 2$	$S_{ij} = \sum_k S_{ijk}$

Diagram illustrating the relationship between observed and true methylation:

- A blue curved arrow points from "Observed methylation" to "True methylation".
- The "True methylation" column shows values  $(p_0, p_1)$ .
- The "Observed methylation" column shows values  $(Y_{ij1}, Y_{ij2}, Y_{ij3}, Y_{ij4})$ .
- The "True methylation" column shows values  $(S_{ij1}, S_{ij2}, S_{ij3}, S_{ij4})$ .

# Model



- ▶ Assume **known error rates**  $p_0$  and  $p_1$ ,

$$\begin{aligned} p_0 &= \mathbb{P}(Y_{ijk} = 1 \mid S_{ijk} = 0) \\ p_1 &= \mathbb{P}(Y_{ijk} = 1 \mid S_{ijk} = 1). \end{aligned} \quad (\text{Lakhal-Chaieb et.al., 2017})$$

- ▶ We specify model

$$\begin{aligned} S_{ij} \mid \mathbf{Z}_i, X_{ij} &\sim \text{Binomial}(X_{ij}, \pi_{ij}) \\ \theta_{ij} = \log \left\{ \frac{\pi_{ij}}{1 - \pi_{ij}} \right\} &= \beta_0(t_{ij}) + \beta_1(t_{ij})Z_{1i} + \beta_2(t_{ij})Z_{2i} + \dots + \beta_P(t_{ij})Z_{Pi}. \end{aligned}$$

- ▶ Use splines to parametrize smooth covariate effects:

$$\beta_p(t_{ij}) = \sum_{l=1}^L \alpha_{pl} B_l(t_{ij}) \text{ for } p = 0, 1, \dots, P.$$

- ▶ Smoothing parameters  $\{\lambda_0, \lambda_1, \dots, \lambda_P\}$  for controlling the smoothness of  $\beta_p(t)$

$$\mathcal{L}^{\text{Penalization}} = \sum_{p=0}^P \lambda_p \int \left( \beta_p''(t) \right)^2 dt = \sum_{p=0}^P \lambda_p \boldsymbol{\alpha}_p^T \mathbf{A} \boldsymbol{\alpha}_p$$

# Methods



## ► Penalized EM algorithm

Initialization:  $\alpha^*$   $\lambda^*$  ;

repeat

    1. E-step: calculate the conditional expected outcomes  $\mathbb{E}(S_{ij} | Y_{ijk})$

    2. M-step:  $(\hat{\alpha}, \hat{\lambda}) = \arg \max Q(\alpha, \lambda | \alpha^*)$ . ( $Q$  is the binomial likelihood replacing  $S_{ij}$  with its expectations)

repeat

    2.1 P-IRLS iteration.

    2.2 Smoothing parameters estimated by REML<sup>‡</sup>.

until estimates converge;

$\hat{\alpha} \rightarrow \alpha^*$  and  $\hat{\lambda} \rightarrow \lambda^*$

until estimates converge;

⇒ Estimates of the smooth functions of covariates effects  
 $\widehat{\beta_1(t)}, \widehat{\beta_2(t)} \dots \widehat{\beta_P(t)}$ .

## ► Inference of smooth covariate effects taking account of the uncertainty in both E step and M step

$$\mathcal{H}(\alpha) = \left\{ \frac{\partial^2 Q(\alpha | \alpha^*)}{\partial \alpha \partial \alpha^T} + \frac{\partial^2 Q(\alpha | \alpha^*)}{\partial \alpha \partial \alpha^{*T}} \right\} \Big|_{\alpha^*=\alpha} ^{\ddagger}$$

⇒ Variance of the smooth estimates  $\widehat{\text{Var}}(\widehat{\beta_p(t)})$

⇒ A region-wise test statistic (p-value) for  $H_0 : \beta_p(t) = 0$

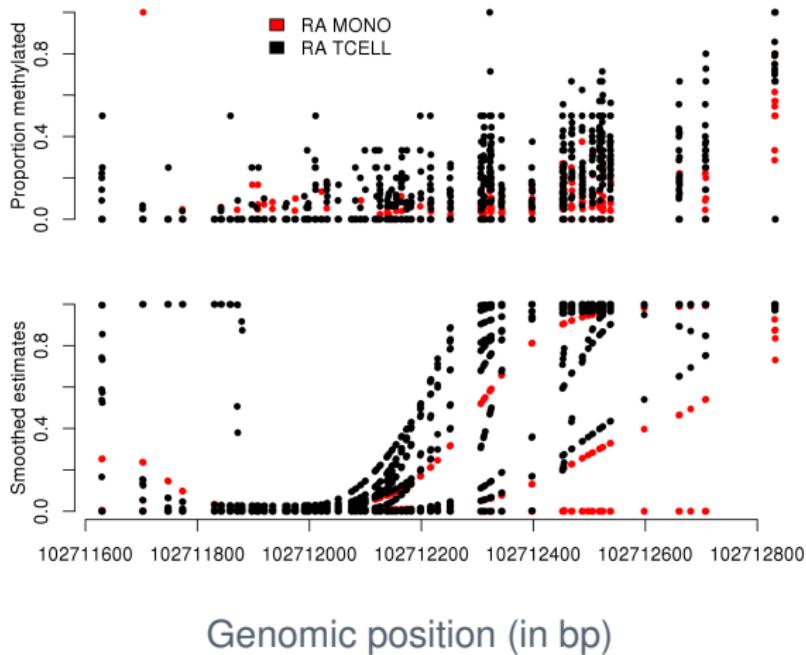
$\widehat{\alpha_p} \left\{ \widehat{\text{Var}}(\widehat{\alpha_p}) \right\}^{-1} \widehat{\alpha_p}^T \sim \chi_{edf}^2$  where

$edf = \text{trace}(2\mathbf{H} - \mathbf{HH}^T)$

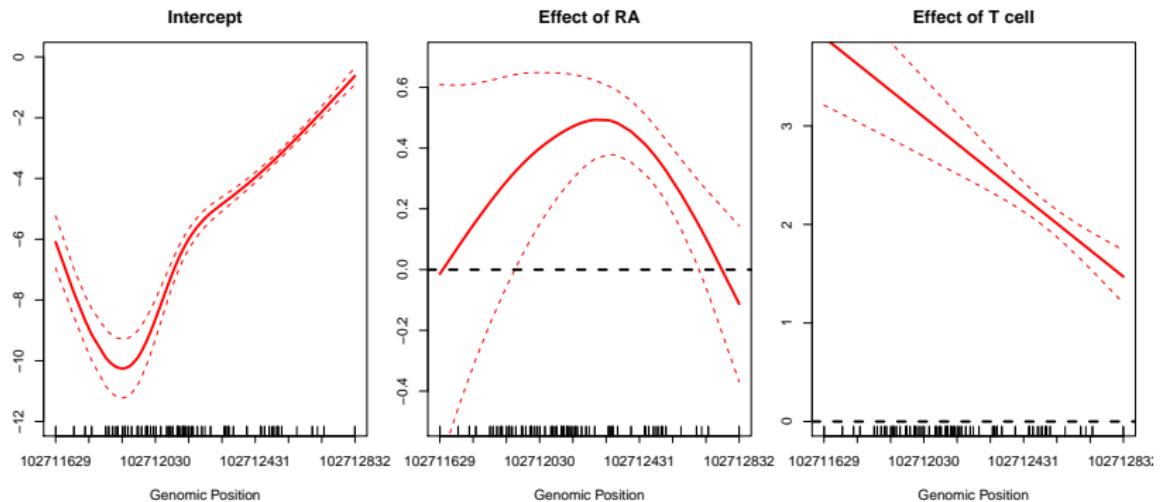
<sup>‡</sup> Oakes, D. (1999) Direct calculation of the information matrix via the EM. JRSSB

# Data example

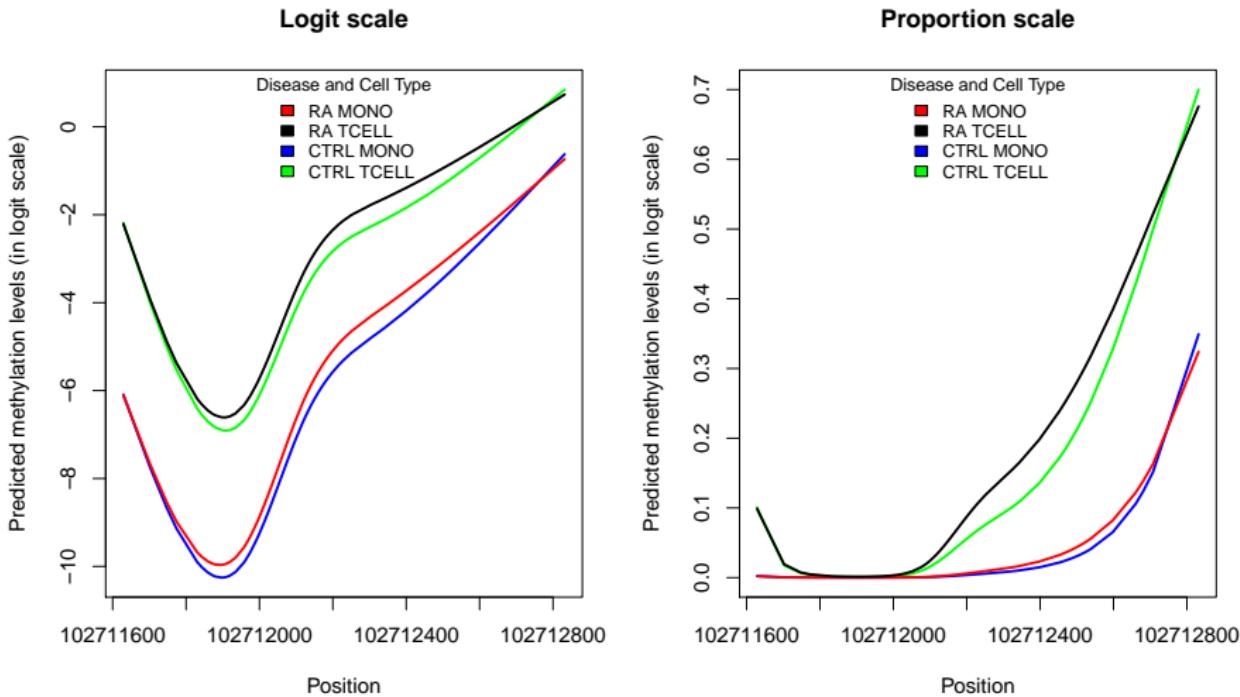
Raw data & per-sample smoothed estimates



# Inference of the smooth covariate effects

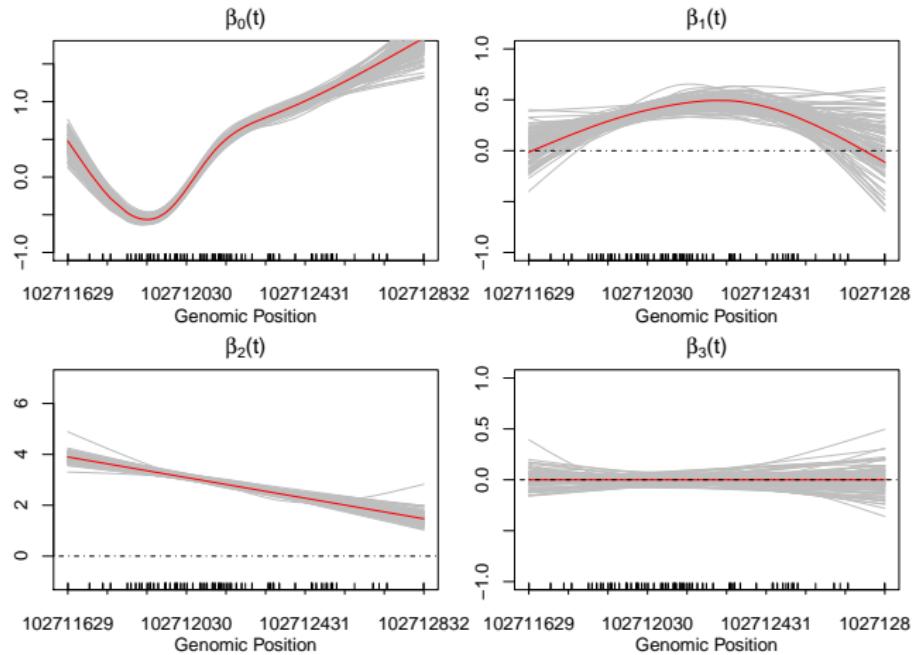


# Predicted methylation levels



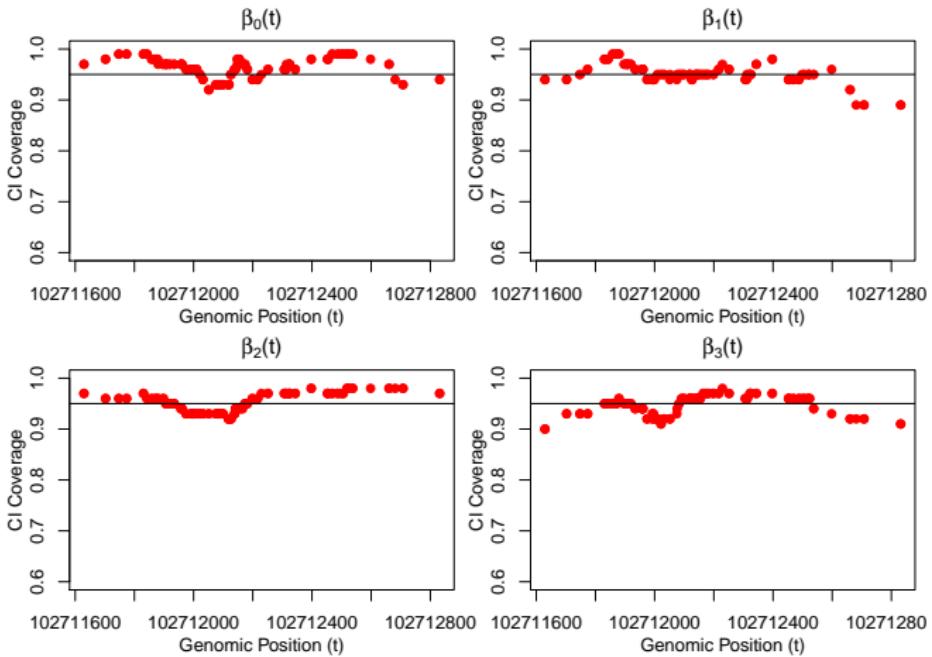
# Simulation

- $\beta_3(t) = 0$  and error parameters  $p_0 = 0.003$  and  $1 - p_1 = 0.1^{\ddagger}$

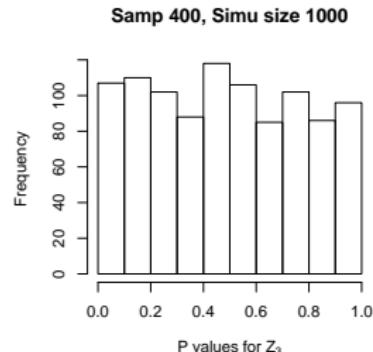
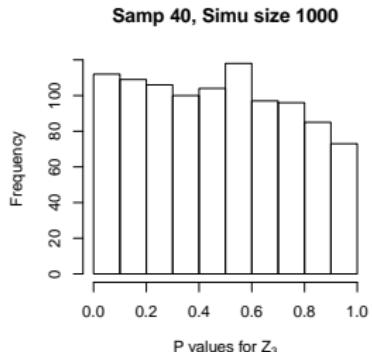
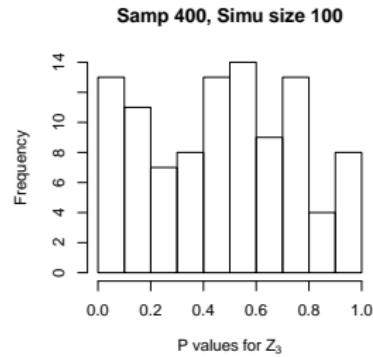
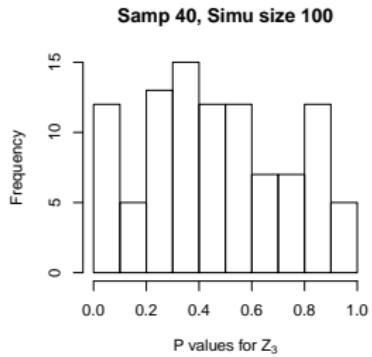


<sup>‡</sup> Prochenka et al. (2015) *Bioinformatics*.

# Coverage probability of confidence intervals



# Region-based p-values under null





- ▶ a **model** characterizing bisulfite sequencing data from multiple samples, which copes naturally with variable read depth, experimental errors and test samples with a mixture of cell types.
- ▶ a **smoothed EM method** to make inference about smooth exposure/covariate effects.
- ▶ the method is shown to be capable in **capturing the major underlying patterns** in the data.
- ▶ **Next step plans**
  - estimating error rate  $p_0$  by calculating C-T conversion rate at non-CpG Cs
  - adding subject-specific random effects to account for the heterogeneity in methylation profiles that cannot be explained by covariates  $Z_s$ s in the model
  - correlated samples

# Acknowledgement

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# Thanks

Questions & Comments