

# Self-Exciting Point Processes for Crime

**Craig Gilmour**  
**Prof. D. Higham**

University of Strathclyde  
*craig.gilmour@strath.ac.uk*

Monday 17th March

# Why Self-exciting Point Processes?

- Studies have shown certain types of crime, e.g. burglaries and gang violence are more likely to occur soon after a previous crime has been committed.
- 'Repeat victimisation' and 'near repeat victimisation' - offences target past victims or occur in nearby locations.
- Self-exciting point processes used in earthquake modelling (e.g. Vere-Jones 1970) - aftershocks occur in the vicinity of the primary earthquake.
- These types of models have been more recently used in the context of crimes (e.g. Mohler 2011).

# Self-exciting Point Processes

Self-exciting point processes can be defined by the *conditional intensity function*

$$\lambda(t) = \lim_{\Delta t \rightarrow 0} \frac{\mathbb{E}(N(t, t + \Delta t) | \mathcal{H}_t)}{\Delta t}.$$

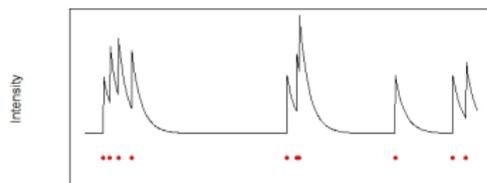
Poisson process

$$\lambda(t) = \mu.$$



Hawkes process (Hawkes 1971)

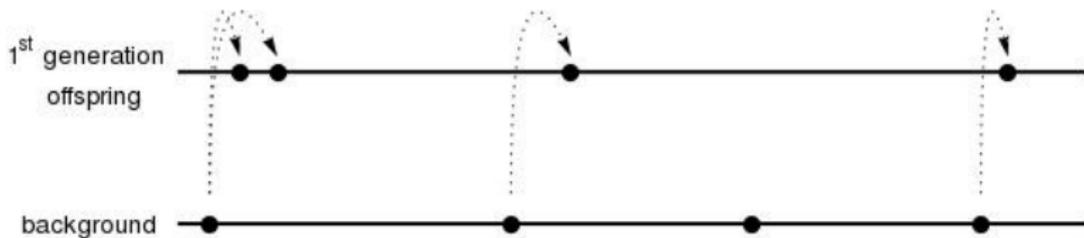
$$\lambda(t) = \mu + \sum_{t_i < t} g(t - t_i).$$



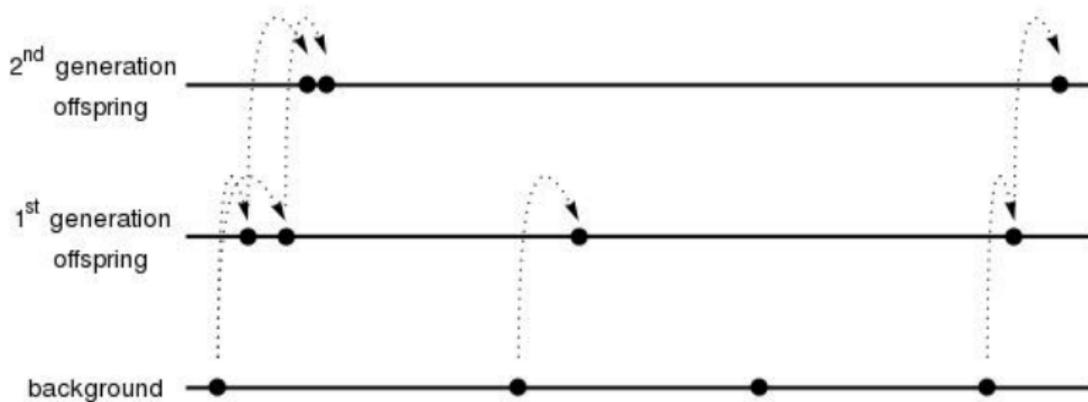
# Self-Exciting Point Process



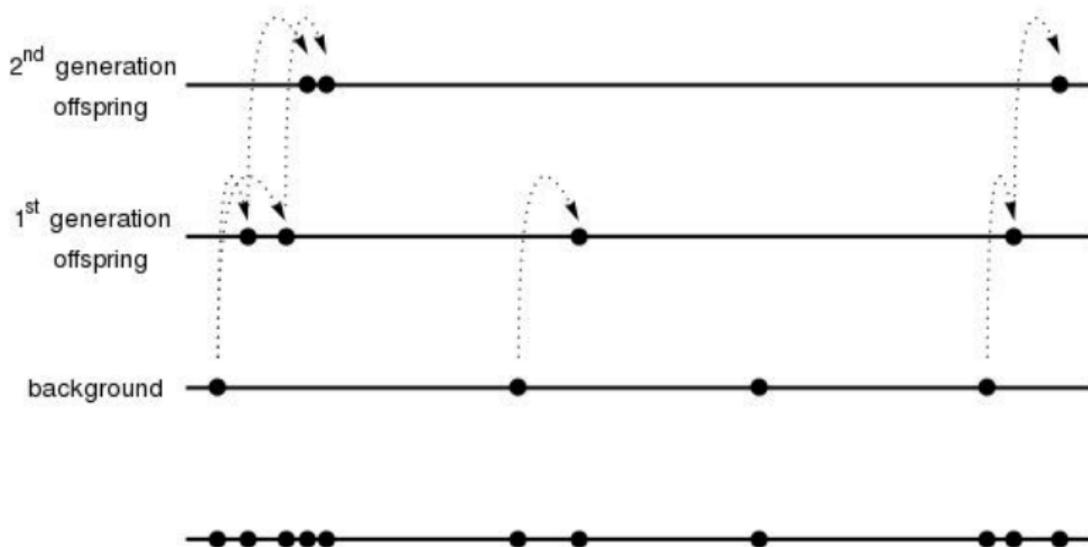
# Self-Exciting Point Process



# Self-Exciting Point Process



# Self-Exciting Point Process



# Parametric Formulation

Often a parametric form is assumed for the triggering function, e.g.

$$g(t - t_i) = \alpha\omega e^{-\omega(t-t_i)}$$

$$\lambda(t) = \mu + \sum_{t_i < t} \alpha\omega e^{-\omega(t-t_i)}$$

Given a point process  $(t_1, t_2, \dots, t_n)$  on an interval  $[0, T)$ , the likelihood function is given by

$$L = \left( \prod_{i=1}^n \lambda(t_i) \right) \exp \left( - \int_0^T \lambda(t) dt \right).$$

For the above ETAS model the log-likelihood can be given as

$$\log L = \sum_{i=1}^n \log \left( \mu + \alpha\omega \sum_{t_j < t_i} e^{-\omega(t_i-t_j)} \right) - \mu T + \sum_{i=1}^n \alpha(e^{-\omega(T-t_i)} - 1).$$

Often the log-likelihood function is very flat, meaning standard iterative methods may not be suitable.

The model can be viewed instead as an incomplete data problem where we don't know whether an event is a background event or has been triggered by a previous event.

For the above ETAS model the log-likelihood can be given as

$$\begin{aligned}\log L = & \sum_{i=1}^n p_{ii} \log(\mu(t_i)) - \int_0^T \mu(t) dt \\ & + \sum_{i=1}^n \sum_{t_j < t_i} p_{ij} \log(g(t_i - t_j)) - \sum_{i=1}^n \int_{t_j}^T g(t - t_i) dt\end{aligned}$$

For a spatio-temporal point process

$$\lambda(x, y, t) = \mu(x, y) + \sum_{t > t_i} g(t - t_i) f(x - x_i, y - y_i)$$

$$g(t) f(x, y) = \alpha \omega e^{-\omega t} \cdot \frac{1}{2\pi\sigma^2} e^{-(x^2 + y^2)/2\sigma^2}$$

We can create the following EM-algorithm

$$\mu = \frac{\sum_{i=1}^n p_{ii}}{T}, \quad \alpha = \frac{\sum_{i=1}^n \sum_{j=1}^{i-1} p_{ij}}{n},$$

$$\omega = \frac{\sum_{i=1}^n \sum_{j=1}^{i-1} p_{ij}}{\sum_{i=1}^n \sum_{j=1}^{i-1} p_{ij} (t_i - t_j)},$$

$$\sigma^2 = \frac{\sum_{i=1}^n \sum_{j=1}^{i-1} p_{ij} ((x_i - x_j)^2 + (y_i - y_j)^2)}{\sum_{i=1}^n \sum_{j=1}^{i-1} 2p_{ij} (t_i - t_j)}.$$

We iterate between this maximisation step, and the following expectation step until convergence

$$p_{ii} = \frac{\mu}{\mu + \sum_{j=1}^{i-1} g(t_i - t_j)},$$
$$p_{ij} = \frac{g(t_i - t_j)}{\mu + \sum_{j=1}^{i-1} g(t_i - t_j)}.$$



# Voronoi Residual Analysis

We can use Voronoi residual analysis to assess the fit of a spatio-temporal point process (Bray 2014).

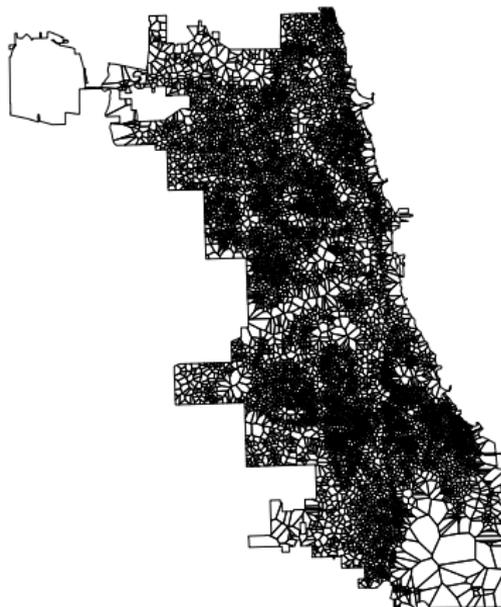
For an event  $(x_i, y_i)$ , the associated Voronoi cell  $C_i$  can be defined as

$$C_i = \{(x, y) \mid d((x, y), (x_i, y_i)) \leq d((x, y), (x_j, y_j)) \quad \forall i \neq j\},$$

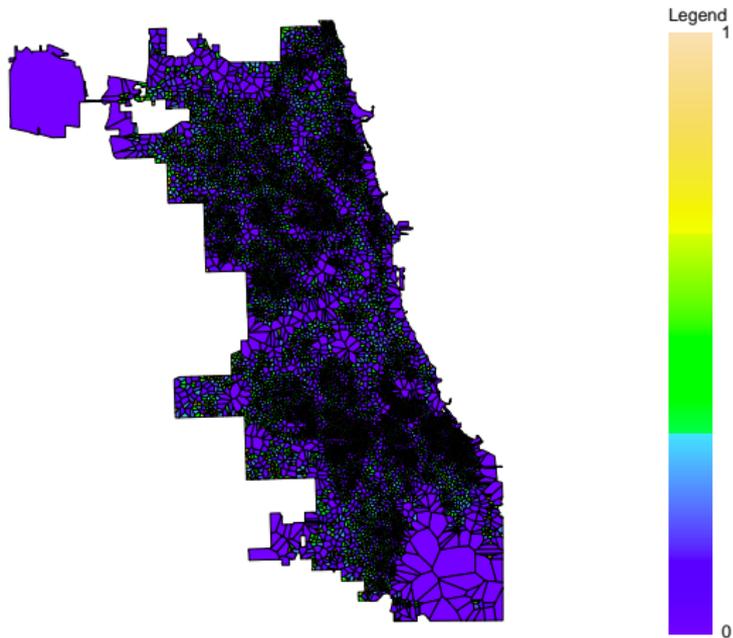
We can evaluate Voronoi residuals  $R_{\text{vor}}^i$  for each Voronoi cell  $i$ , where

$$R_{\text{vor}}^i = 1 - \int_0^T \int \int_{(x,y) \in C_i} \lambda(x, y, t) dx dy dt.$$

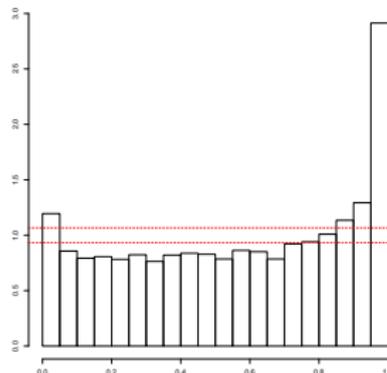
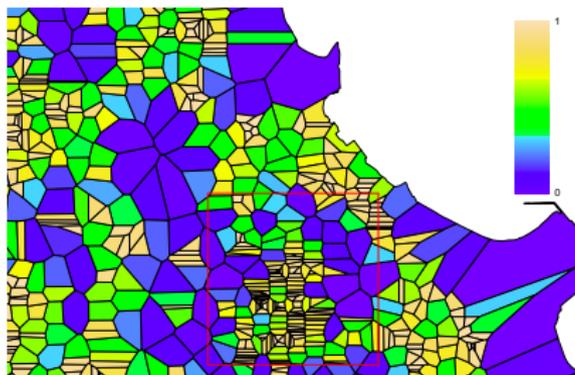
# Voronoi Residual Analysis



# Voronoi Residual Analysis



# Voronoi Residual Analysis



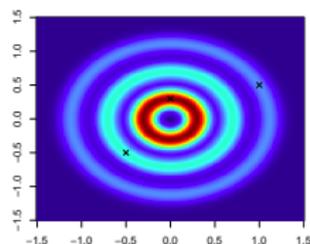
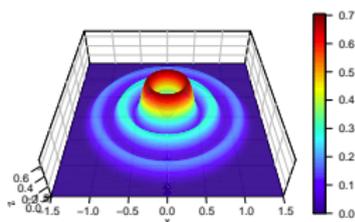
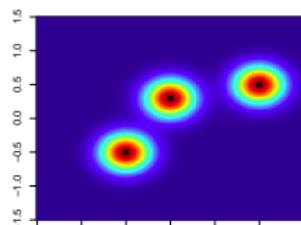
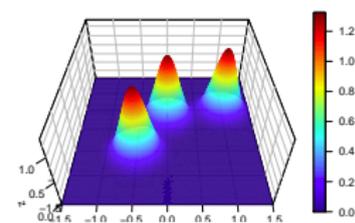
Modelling crime with a homogeneous background rate is not suitable.

Marsan and Lengliné (2008) introduced Model Independent Stochastic Declustering (MISD) in order to find a non-parametric form for the triggering function without any prior assumptions.

This has been used to model crime events (e.g. Mohler et al. 2011).

We use different adaptations of this to model crime in Chicago.

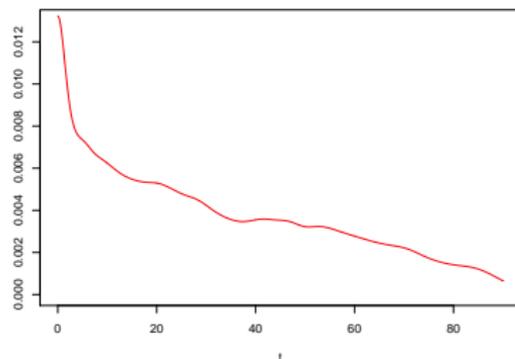
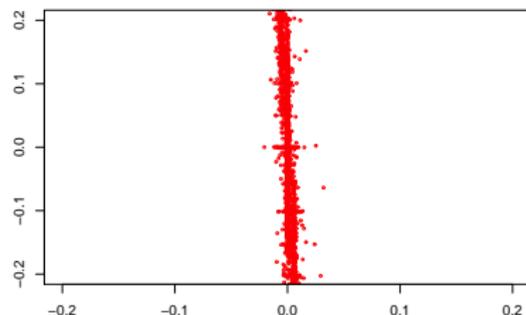
We model the crime assuming isotropy in the triggering component.



# Results on Chicago Data

Results using model from Mohler et al. 2011.

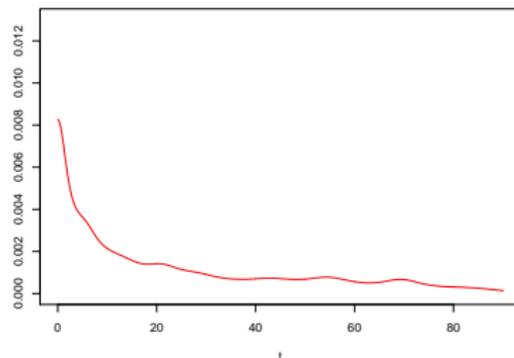
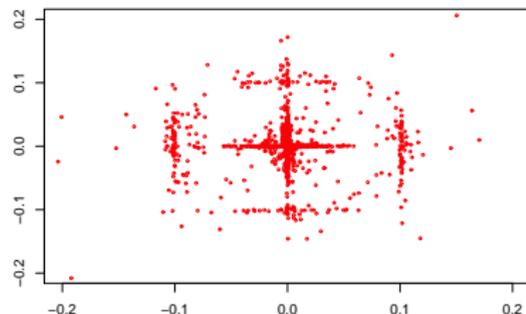
$$\lambda(x, y, t | \mathcal{H}_t) = \nu(t)\mu(x, y) + \sum_{t > t_i} g(x - x_i, y - y_i, t - t_i).$$



# Results on Chicago Data

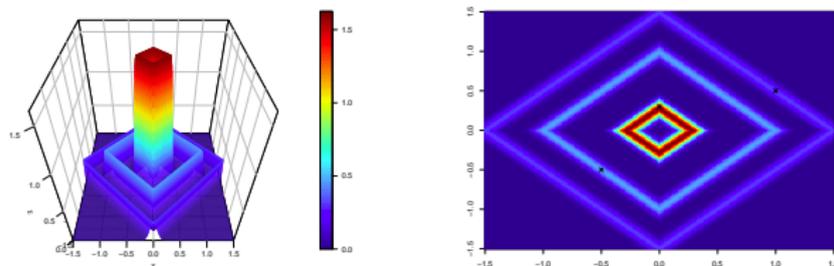
Results when assuming isotropy.

$$\lambda(x, y, t | \mathcal{H}_t) = \nu(t)\mu(x, y) + \sum_{t > t_i} g(r - r_i, t - t_i).$$



# Non-parametric Methods

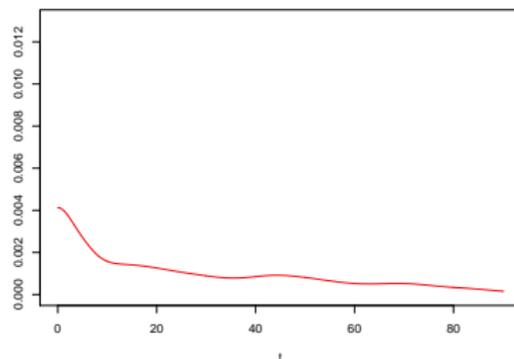
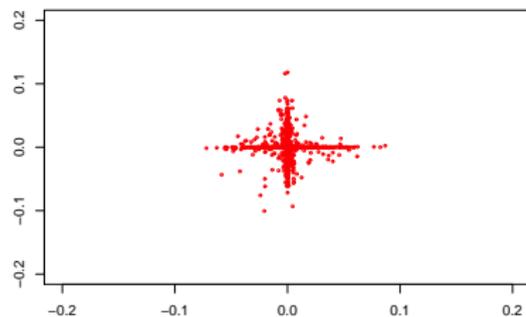
We also model the triggering function using Manhattan distance to measure the distance between events.



# Results on Chicago Data

Results when using Manhattan distance.

$$\lambda(x, y, t | \mathcal{H}_t) = \nu(t)\mu(x, y) + \sum_{t > t_i} g(d - d_i, t - t_i).$$



- Tested self-exciting point process models on Chicago crime data
- Tested variants of the standard approaches
- Could maintain predictive power with reduced dimensionality

# The End