# Dimension reduction in nonlinear Bayesian inverse problems 

Youssef Marzouk<br>joint work with Daniele Bigoni, Michael Brennan, Tiangang Cui, Kody Law, Alessio Spantini, Olivier Zahm<br>Department of Aeronautics and Astronautics<br>Center for Computational Science and Engineering<br>Statistics and Data Science Center<br>Massachusetts Institute of Technology<br>http://uqgroup.mit.edu<br>Support from AFOSR, DOE, NSF, ONR

2 November 2021

## Motivation: inverse problems in the Bayesian setting

Observations y
Parameters x


$$
\pi_{\mathrm{pos}}(x):=\underbrace{\pi(x \mid y) \propto \mathcal{L}_{y}(x) \pi_{\mathrm{pr}}(x)}_{\text {Bayes' rule }}
$$

- Characterize the posterior distribution (density $\pi_{\text {pos }}$ )
- This is a challenging task since:
- $x \in \mathbb{R}^{d}$ is typically high-dimensional (e.g., a discretized function)
- $\pi_{\text {pos }}$ is non-Gaussian
- evaluations of the likelihood (hence $\pi_{\text {pos }}$ ) may be expensive
- $\pi_{\text {pos }}$ can be evaluated up to a normalizing constant


## A conjecture

In many situations, the data are "informative" only on a low-dimensional subspace


$$
\mathbb{R}^{d}=\underbrace{X_{r}}_{\pi_{\mathrm{pos}} \neq \pi_{\mathrm{pr}}}+\underbrace{X_{\perp}}_{\pi_{\mathrm{pos}} \approx \pi_{\mathrm{pr}}}
$$

## A conjecture

In many situations, the data are "informative" only on a low-dimensional subspace


This structure is now well understood in the linear-Gaussian case, $x \sim N\left(0, \Sigma_{\text {pr }}\right), y \mid x \sim N\left(G x, \Sigma_{\text {obs }}\right)$ [Spantini et al. 2015]:

- Optimal approximations of the posterior covariance as a low-rank update of the prior, $\widetilde{\Sigma}_{\text {pos }}=\Sigma_{\text {pr }}-K_{r} K_{r}^{\top}$, for any $r \leq d$
- Optimal posterior mean approximations, $\widetilde{\mu}_{\text {pos }}=A_{r} y$
- Central role of generalized eigenproblems, e.g., $\left(G^{T} \Sigma_{\mathrm{obs}}^{-1} G, \Sigma_{\mathrm{pr}}^{-1}\right)$


## Low effective dimensionality of Bayesian inverse problems

More general idea: the posterior distribution can be well approximated by

$$
\widetilde{\pi}_{\mathrm{pos}}(x) \propto \widetilde{\mathcal{L}}\left(P_{r} x\right) \pi_{\mathrm{pr}}(x)
$$

for some positive function $\widetilde{\mathcal{L}}$ and rank $r$ linear projector $P_{r} \in \mathbb{R}^{d \times d}$

## Low effective dimensionality of Bayesian inverse problems

More general idea: the posterior distribution can be well approximated by

$$
\widetilde{\pi}_{\mathrm{pos}}(x) \propto \widetilde{\mathcal{L}}\left(P_{r} x\right) \pi_{\mathrm{pr}}(x)
$$

for some positive function $\widetilde{\mathcal{L}}$ and rank $r$ linear projector $P_{r} \in \mathbb{R}^{d \times d}$

## induces a decomposition of the space

$$
x=x_{r}+x_{\perp} \quad \begin{cases}x_{r} & \in \operatorname{Im}\left(P_{r}\right) \\ x_{\perp} & \in \operatorname{Ker}\left(P_{r}\right)\end{cases}
$$

By construction, $x \mapsto \widetilde{\mathcal{L}}\left(P_{r} x\right)=\widetilde{\mathcal{L}}\left(x_{r}\right)$ is only a function of $x_{r} \in \operatorname{Im}\left(P_{r}\right) \equiv \mathbb{R}^{r}$.

## Low effective dimensionality of Bayesian inverse problems

More general idea: the posterior distribution can be well approximated by

$$
\widetilde{\pi}_{\mathrm{pos}}(x) \propto \widetilde{\mathcal{L}}\left(P_{r} x\right) \pi_{\mathrm{pr}}(x)
$$

for some positive function $\widetilde{\mathcal{L}}$ and rank $r$ linear projector $P_{r} \in \mathbb{R}^{d \times d}$

## induces a decomposition of the space

$$
x=x_{r}+x_{\perp} \quad \begin{cases}x_{r} & \in \operatorname{Im}\left(P_{r}\right) \\ x_{\perp} & \in \operatorname{Ker}\left(P_{r}\right)\end{cases}
$$

By construction, $x \mapsto \widetilde{\mathcal{L}}\left(P_{r} x\right)=\widetilde{\mathcal{L}}\left(x_{r}\right)$ is only a function of $x_{r} \in \operatorname{Im}\left(P_{r}\right) \equiv \mathbb{R}^{r}$. If $r \ll d$, we can:

- Design structure-exploiting MCMC algorithms to sample from $\pi_{\text {pos }}$ (e.g., DILI samplers [Cui, Law, M 2016])
- More easily build surrogates (i.e., fast approximations) of $x_{r} \mapsto \widetilde{\mathcal{L}}\left(x_{r}\right)$
- Develop tractable variational characterizations of the posterior (second part of this talk)


## Many previous proposals

- $P_{r}$ can be defined as a projector onto the dominant eigenspace of a matrix
$\mathbf{H} \in \mathbb{R}^{d \times d}$ which contains "relevant information"


## Many previous proposals

- $P_{r}$ can be defined as a projector onto the dominant eigenspace of a matrix $\mathbf{H} \in \mathbb{R}^{d \times d}$ which contains "relevant information"
- Likelihood-informed subspace (LIS) [Cui et al. 2014]

$$
\mathbf{H}_{\mathrm{LIS}}=\int(\nabla G)^{T} \Sigma_{\mathrm{obS}}^{-1}(\nabla G) \mathrm{d} \pi_{\mathrm{pos}}
$$

where $\mathcal{L}_{y}$ follows from $y \sim \mathcal{N}\left(G(x), \Sigma_{\text {obs }}\right)$

- Active subspace (AS) [Constantine et al. 2015]

$$
\mathbf{H}_{\mathrm{AS}}=\int \nabla \log \mathcal{L}_{y} \otimes \nabla \log \mathcal{L}_{y} \mathrm{~d} \pi_{\mathrm{pr}}
$$

## Many previous proposals

- $P_{r}$ can be defined as a projector onto the dominant eigenspace of a matrix $\mathbf{H} \in \mathbb{R}^{d \times d}$ which contains "relevant information"
- Likelihood-informed subspace (LIS) [Cui et al. 2014]

$$
\mathbf{H}_{\mathrm{LIS}}=\int(\nabla G)^{T} \Sigma_{\mathrm{obS}}^{-1}(\nabla G) \mathrm{d} \pi_{\mathrm{pos}}
$$

where $\mathcal{L}_{y}$ follows from $y \sim \mathcal{N}\left(G(x), \Sigma_{\text {obs }}\right)$

- Active subspace (AS) [Constantine et al. 2015]

$$
\mathbf{H}_{\mathrm{AS}}=\int \nabla \log \mathcal{L}_{y} \otimes \nabla \log \mathcal{L}_{y} \mathrm{~d} \pi_{\mathrm{pr}}
$$

- Similarly, various definitions of $\widetilde{\mathcal{L}}$ :
- (LIS) Fix complementary parameters $\widetilde{\mathcal{L}}\left(P_{r} x\right)=\mathcal{L}_{y}\left(P_{r} x+\left(I-P_{r}\right) m_{0}\right)$
- (AS) Take conditional expectation of the log-likelihood

$$
\widetilde{\mathcal{L}}\left(P_{r} x\right)=\exp \mathbb{E}_{\pi_{\mathrm{rr}}}\left(\log \mathcal{L}_{y} \mid P_{r} x\right)
$$

## Broad objective

Build an approximation of $\pi_{\text {pos }}$ of the form

$$
\widetilde{\pi}_{\mathrm{pos}}(x) \propto \widetilde{\mathcal{L}}\left(P_{r} x\right) \pi_{\mathrm{pr}}(x) \quad \text { with }\left\{\begin{array}{l}
\widetilde{\mathcal{L}}: \mathbb{R}^{d} \rightarrow \mathbb{R}^{+} \\
P_{r} \in \mathbb{R}^{d \times d} \text { rank- } r \text { projector }
\end{array}\right.
$$

such that

$$
D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}} \| \widetilde{\pi}_{\mathrm{pos}}\right) \leq \varepsilon
$$

with $r=r(\varepsilon)$ much smaller than $d$.
See full details in [ZCLSM 21].

## Decomposition of the error

## A "Pythagorean" theorem

For any $P_{r}$ and $\widetilde{\mathcal{L}}$ we have

$$
D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}} \| \widetilde{\pi}_{\mathrm{pos}}\right)=\underbrace{D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}} \| \pi_{\mathrm{pos}}^{*}\right)}_{=\text {function }\left(P_{r}\right)}+\underbrace{D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}}^{*} \| \widetilde{\pi}_{\mathrm{pos}}\right)}_{=\text {function }\left(P_{r}, \tilde{\mathcal{L}}\right)}
$$

where

$$
\pi_{\mathrm{pos}}^{*}(x) \propto \mathbb{E}_{\pi_{\mathrm{pr}}}\left(\mathcal{L}_{y} \mid P_{r} x\right) \pi_{\mathrm{pr}}(x)
$$

## Decomposition of the error

## A "Pythagorean" theorem

For any $P_{r}$ and $\widetilde{\mathcal{L}}$ we have

$$
D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}} \| \tilde{\pi}_{\mathrm{pos}}\right)=\underbrace{D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}} \| \pi_{\mathrm{pos}}^{*}\right)}_{=\text {function }\left(P_{r}\right)}+\underbrace{D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}}^{*}| | \tilde{\pi}_{\mathrm{pos}}\right)}_{=\text {function }\left(P_{r}, \widetilde{\mathcal{L}}\right)}
$$

where

$$
\pi_{\mathrm{pos}}^{*}(x) \propto \mathbb{E}_{\pi_{\mathrm{pr}}}\left(\mathcal{L}_{y} \mid P_{r} x\right) \pi_{\mathrm{pr}}(x)
$$

This allows decoupling the construction of $\widetilde{\mathcal{L}}$ and $P_{r}$.

- Given $P_{r}$, the function $\widetilde{\mathcal{L}}$ such that $\widetilde{\mathcal{L}}\left(P_{r} x\right)=\mathbb{E}_{\pi_{\mathrm{pr}}}\left(\mathcal{L}_{y} \mid P_{r} x\right)$ yields

$$
D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}}^{*} \| \widetilde{\pi}_{\mathrm{pos}}\right)=0
$$

- How to construct $P_{r}$ such that

$$
D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}} \| \pi_{\mathrm{pos}}^{*}\right) \leq \varepsilon
$$

with a rank $r \ll d$ ?

## Constructing the projector $P_{r}$

## Assumption on the prior distribution

There exist functions $V$ and $\Psi$ such that

$$
\pi_{\mathrm{pr}}(x) \propto \exp (-V(x)-\Psi(x)) \quad \text { with } \quad\left\{\begin{array}{l}
\nabla^{2} V \succeq \Gamma \\
\exp (\sup \Psi-\inf \Psi) \leq \kappa
\end{array}\right.
$$

for some SPD matrix $\Gamma \in \mathbb{R}^{d \times d}$ and some $\kappa \geq 1$.

## Constructing the projector $P_{r}$

## Assumption on the prior distribution

There exist functions $V$ and $\Psi$ such that

$$
\pi_{\mathrm{pr}}(x) \propto \exp (-V(x)-\Psi(x)) \quad \text { with } \quad\left\{\begin{array}{l}
\nabla^{2} V \succeq \Gamma \\
\exp (\sup \Psi-\inf \Psi) \leq \kappa
\end{array}\right.
$$

for some SPD matrix $\Gamma \in \mathbb{R}^{d \times d}$ and some $\kappa \geq 1$.


- Gaussian prior satisfies this assumption with $\Gamma=\Sigma_{\text {pr }}^{-1}$ and $\kappa=1$
- Gaussian mixture $\pi_{\mathrm{pr}} \propto \sum_{i} \mathcal{N}\left(\mu_{i}, \Sigma_{i}\right)$ also satisfies this assumption
- Uniform prior on convex bounded domain also allowed [ZCLSM21]


## Constructing the projector $P_{r}$

Based on this assumption, $\pi_{\mathrm{pr}}$ satisfies the logarithmic Sobolev inequality

$$
\int h^{2} \log \frac{h^{2}}{\int h^{2} \mathrm{~d} \pi_{\mathrm{pr}}} \mathrm{~d} \pi_{\mathrm{pr}} \leq 2 \kappa \int\|\nabla h\|_{\Gamma_{-1}^{2}}^{2} \mathrm{~d} \pi_{\mathrm{pr}}
$$

for any function $h$ with sufficient regularity.

- Putting $h^{2}=\mathcal{L}_{y} / \int \mathcal{L}_{y} \mathrm{~d} \pi_{\mathrm{pr}}$ bounds the KL divergence from prior to posterior:

$$
\mathcal{D}_{\mathrm{KL}}\left(\pi_{\mathrm{pos}} \| \pi_{\mathrm{pr}}\right) \leq \frac{\kappa}{2} \int\left\|\nabla \log \mathcal{L}_{y}\right\|_{\Gamma^{-1}}^{2} \mathrm{~d} \pi_{\mathrm{pos}}
$$

## Constructing the projector $P_{r}$

## Proposition: subspace logarithmic Sobolev inequality

$\pi_{\mathrm{pr}}$ also satisfies

$$
\int h^{2} \log \frac{h^{2}}{\mathbb{E}\left(h^{2} \mid P_{r} x\right)} \mathrm{d} \pi_{\mathrm{pr}} \leq 2 \kappa \int\left\|\left(I_{d}-P_{r}^{T}\right) \nabla h\right\|_{\Gamma^{-1}}^{2} \mathrm{~d} \pi_{\mathrm{pr}}
$$

for any function $h$ with sufficient regularity and any projector $P_{r}$.

## Constructing the projector $P_{r}$

## Proposition: subspace logarithmic Sobolev inequality

$\pi_{\text {pr }}$ also satisfies

$$
\int h^{2} \log \frac{h^{2}}{\mathbb{E}\left(h^{2} \mid P_{r} x\right)} \mathrm{d} \pi_{\mathrm{pr}} \leq 2 \kappa \int\left\|\left(I_{d}-P_{r}^{T}\right) \nabla h\right\|_{r^{-1}}^{2} \mathrm{~d} \pi_{\mathrm{pr}}
$$

for any function $h$ with sufficient regularity and any projector $P_{r}$.

## Corollary

For any projector $P_{r}$ we have

$$
D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}} \| \pi_{\mathrm{pos}}^{*}\right) \leq \frac{\kappa}{2} \mathcal{R}_{\pi_{\mathrm{pos}}}\left(P_{r}\right)
$$

where

$$
\mathcal{R}_{\pi_{\mathrm{pos}}}\left(P_{r}\right)=\int\left\|\left(I_{d}-P_{r}^{T}\right) \nabla \log \mathcal{L}_{y}\right\|_{\Gamma-1}^{2} \mathrm{~d} \pi_{\mathrm{pos}}
$$

## Constructing the projector $P_{r}$

## Corollary

For any projector $P_{r}$ we have

$$
D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}} \| \pi_{\mathrm{pos}}^{*}\right) \leq \frac{\kappa}{2} \mathcal{R}_{\pi_{\mathrm{pos}}}\left(P_{r}\right)
$$

where

$$
\mathcal{R}_{\pi_{\text {pos }}}\left(P_{r}\right)=\int\left\|\left(I_{d}-P_{r}^{T}\right) \nabla \log \mathcal{L}_{y}\right\|_{\Gamma-1}^{2} \mathrm{~d} \pi_{\text {pos }}
$$

## Constructing the projector $P_{r}$

## Corollary

For any projector $P_{r}$ we have

$$
D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}} \| \pi_{\mathrm{pos}}^{*}\right) \leq \frac{\kappa}{2} \mathcal{R}_{\pi_{\mathrm{pos}}}\left(P_{r}\right)
$$

where

$$
\mathcal{R}_{\pi_{\text {pos }}}\left(P_{r}\right)=\int\left\|\left(I_{d}-P_{r}^{T}\right) \nabla \log \mathcal{L}_{y}\right\|_{\Gamma_{-1}}^{2} \mathrm{~d} \pi_{\mathrm{pos}}
$$

Finding $P_{r}$ that minimizes this bound corresponds to PCA of $\nabla \log \mathcal{L}_{y}(X)$.

- For a fixed $r$, the minimizer $P_{r}^{*}$ of the reconstruction error $\mathcal{R}_{\pi_{\text {pos }}}\left(P_{r}\right)$ is the $\Gamma$-orthogonal projector onto the dominant generalized eigenspace of

$$
\mathbf{H}=\int \nabla \log \mathcal{L}_{y} \otimes \nabla \log \mathcal{L}_{y} \mathrm{~d} \pi_{\mathrm{pos}}
$$

- Furthermore, we have $\mathcal{R}_{\pi_{\mathrm{pos}}}\left(P_{r}^{*}\right)=\sum_{i>r} \lambda_{i}$, where $\lambda_{i}$ is the $i$-th generalized eigenvalue of $(\mathbf{H}, \Gamma)$


## An idealized algorithm

1 Compute

$$
\mathbf{H}=\int \nabla \log \mathcal{L}_{y} \otimes \nabla \log \mathcal{L}_{y} \mathrm{~d} \pi_{\mathrm{pos}}
$$

2 Define $P_{r}$ as the projector on the dominant eigenspace of $\mathbf{H}$
3 Compute the conditional expectation

$$
\widetilde{\mathcal{L}}\left(P_{r} x\right)=\mathbb{E}_{\mathrm{pr}}\left(\mathcal{L}_{y} \mid P_{r} x\right)
$$

Then $\pi_{\text {pos }}^{*}(x) \propto \widetilde{\mathcal{L}}\left(P_{r} x\right) \pi_{\text {pr }}(x)$ satisfies

$$
D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}} \| \pi_{\mathrm{pos}}^{*}\right) \leq \frac{\kappa}{2} \sum_{i>r} \lambda_{i}
$$

- At step 2, we can choose the rank $r=r(\varepsilon)$ of $P_{r}$ such that

$$
D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}} \| \pi_{\mathrm{pos}}^{*}\right) \leq \varepsilon
$$

- A strong decay in $\lambda_{i}$ implies $r(\varepsilon) \ll d$


## An idealized algorithm

1 Compute

$$
\mathbf{H}=\int \nabla \log \mathcal{L}_{y} \otimes \nabla \log \mathcal{L}_{y} \mathrm{~d} \pi_{\mathrm{pos}}
$$

2 Define $P_{r}$ as the projector on the dominant eigenspace of $\mathbf{H}$
3 Compute the conditional expectation

$$
\widetilde{\mathcal{L}}\left(P_{r} x\right)=\mathbb{E}_{\mathrm{pr}}\left(\mathcal{L}_{y} \mid P_{r} x\right)
$$

## Practical issues

- Evaluating $\mathbf{H}$ requires computing an integral over the posterior
- Computing the conditional expectation requires some effort


## Sample approximations

(1) Monte Carlo approximation of $\mathbf{H}$ :

$$
\mathbf{H} \approx \widehat{\mathbf{H}}_{K}:=\frac{1}{K} \sum_{i=1}^{K} \nabla \log \mathcal{L}_{y}\left(X_{i}\right) \otimes \nabla \log \mathcal{L}_{y}\left(X_{i}\right) \quad \text { with } \quad X_{i} \stackrel{\mathrm{iid}}{\sim} \pi_{\text {pos }}
$$

## Proposition

Under some assumptions, quasi-optimal projectors are obtained with high probability $1-\delta$ if

$$
K \geq \mathcal{O}\left(\sqrt{\operatorname{rank}(H)}+\sqrt{\log \left(2 \delta^{-1}\right)}\right)^{2}
$$

- Key assumption: $\nabla \log \mathcal{L}_{y}(X)$ is sub-Gaussian, for $X \sim \pi_{\text {pos }}$
(2) Sample approximations of the conditional expectation $\mathbb{E}_{\mathrm{pr}}\left(\mathcal{L}_{y} \mid P_{r} x\right)$
- Error controlled by same factors; details in [ZCLSM21]

Example: GOMOS atmospheric remote sensing [Tamminen 2004]

- Estimate gas densities $x=\varrho^{\text {gas }}(z)$ from transmission spectra $y_{\omega}(z)$
- Beer's law:

$$
y_{\omega}(z)=\exp \left(-\int_{\text {light path }} \sum_{\text {gas }} \alpha_{\omega}^{\text {gas }}(z(\zeta)) \varrho^{\text {gas }}(z(\zeta)) \mathrm{d} \zeta\right)+\xi
$$

Signals:


- Gaussian prior $\mathcal{N}\left(\mu_{\mathrm{pr}}, \Sigma_{\mathrm{pr}}\right)$ (hence $\Gamma=\Sigma_{\mathrm{pr}}^{-1}$ and $\kappa=1$ )
- After discretization of the atmosphere, $\operatorname{dim}(x)=200$


## GOMOS: results



## GOMOS: results



## GOMOS: results

$$
\begin{aligned}
& D_{\mathrm{KL}}\left(\pi_{\text {pos }} \| \widetilde{\pi}_{\text {pos }}\right)=\text { function }(r) \\
& \mathbf{H}^{(\rho)}=\int \nabla \log \mathcal{L}_{y} \otimes \nabla \log \mathcal{L}_{y} \mathrm{~d} \rho \\
& \mathbf{H}_{\mathrm{LIS}}^{(\rho)}=\int(\nabla G)^{T} \Gamma_{\text {obs }}^{-1}(\nabla G) \mathrm{d} \rho
\end{aligned}
$$

## GOMOS: results



## An iterative algorithm

In practice, to avoid drawing samples from $\pi_{\text {pos }}$, we can iterate directly towards a low-dimensional approximation $\widetilde{\pi}_{\text {pos }}$ :

## Conceptually:

$$
\left(\rho^{\ell} \equiv \widetilde{\pi}_{\mathrm{pos}}^{r, \ell}\right) \xrightarrow{\text { sampling }} H^{\left(\rho^{\ell+1}\right)} \xrightarrow{\text { eigenprob }} P_{r}^{\ell+1} \longrightarrow\left(\rho^{\ell+1} \equiv \widetilde{\pi}_{\mathrm{pos}}^{r, \ell+1}\right) \rightarrow \cdots
$$

## Iterative algorithm: results


(left) fixed threshold; (right) fixed rank

## Questions about these low-dimensional approximations

Some open or interesting questions:

- Many MCMC algorithms use the subspace $\operatorname{Im}\left(P_{r}\right)$ to derive proposals and/or splitting (Metropolis-within-Gibbs) schemes (e.g., DILI [Cui et al. 2016])
- Impact of subspace quality on computational performance of MCMC algorithms? Some inital results in [Cui \& Tong 2021]
- Understanding the convergence of iterative algorithms for identifying the projector $P_{r}$, and the associated computational tradeoffs
- Extension to the infinite-dimensional setting
- Possibility of handling heavier-tailed priors?


## Questions about these low-dimensional approximations

Some open or interesting questions:

- Many MCMC algorithms use the subspace $\operatorname{Im}\left(P_{r}\right)$ to derive proposals and/or splitting (Metropolis-within-Gibbs) schemes (e.g., DILI [Cui et al. 2016])
- Impact of subspace quality on computational performance of MCMC algorithms? Some inital results in [Cui \& Tong 2021]
- Understanding the convergence of iterative algorithms for identifying the projector $P_{r}$, and the associated computational tradeoffs
- Extension to the infinite-dimensional setting
- Possibility of handling heavier-tailed priors?

Next: an application of these ideas to transport. . .

## Inference through transport

Main idea: Characterize $\pi_{\text {pos }}$ (henceforth $\pi$ ) as a transformation of some simple distribution $\rho$.

Goal: Find a function $T$ s.t. if $X \sim \rho$, then $T(X) \sim \pi$.


## Inference through transport

Main idea: Characterize $\pi_{\text {pos }}$ (henceforth $\pi$ ) as a transformation of some simple distribution $\rho$.

Goal: Find a function $T$ s.t. if $X \sim \rho$, then $T(X) \sim \pi$.


## Inference through transport

Main idea: Characterize $\pi_{\text {pos }}$ (henceforth $\pi$ ) as a transformation of some simple distribution $\rho$.

Goal: Find a function $T$ s.t. if $X \sim \rho$, then $T(X) \sim \pi$.


## Inference through transport

Main idea: Characterize $\pi_{\text {pos }}$ (henceforth $\pi$ ) as a transformation of some simple distribution $\rho$.

Goal: Find a function $T$ s.t. if $X \sim \rho$, then $T(X) \sim \pi$.


Notation: $\begin{gathered}\text { pushforward } \\ T_{\sharp} \rho=\pi\end{gathered} \begin{gathered}\text { pullback } \\ \sharp\end{gathered} T^{\sharp} \pi$

## How to construct a suitable map?

Maps from unnormalized densities, i.e., variational characterization of the map $T$ :

## How to construct a suitable map?

Maps from unnormalized densities, i.e., variational characterization of the map $T$ :

$$
\min _{T \in \mathcal{T}^{h}} \mathcal{D}_{K L}\left(T_{\sharp} \rho \| \pi\right)=\min _{T \in \mathcal{T}^{h}} \mathcal{D}_{K L}\left(\rho \| T_{\sharp}^{-1} \pi\right)
$$

$-\pi$ is the "target" density on $\mathbb{R}^{d} ; \rho$ is, e.g., $\mathcal{N}\left(0, \mathbf{I}_{d}\right)$

- $\mathcal{T}^{h}$ is a parameterized class of maps from $\mathbb{R}^{d}$ to itself
- For instance, monotone lower triangular maps (approximate the Knothe-Rosenblatt rearrangement)
- Expectation is with respect to the reference measure $\rho$
- Compute via, e.g., Monte Carlo, sparse quadrature
- Use unnormalized evaluations of $\pi$ and its gradients
- No MCMC or importance sampling
- In general non-convex, unless $\pi$ is log-concave


## How to construct a suitable map?

Maps from unnormalized densities, i.e., variational characterization of the map $T$ :

$$
\min _{T \in \mathcal{T}^{h}} \mathcal{D}_{K L}\left(T_{\sharp} \rho \| \pi\right)=\min _{T \in \mathcal{T}^{h}} \mathcal{D}_{K L}\left(\rho \| T_{\sharp}^{-1} \pi\right)
$$

$-\pi$ is the "target" density on $\mathbb{R}^{d} ; \rho$ is, e.g., $\mathcal{N}\left(0, \mathbf{I}_{d}\right)$

- $\mathcal{T}^{h}$ is a parameterized class of maps from $\mathbb{R}^{d}$ to itself
- For instance, monotone lower triangular maps (approximate the Knothe-Rosenblatt rearrangement)
- Expectation is with respect to the reference measure $\rho$
- Compute via, e.g., Monte Carlo, sparse quadrature
- Use unnormalized evaluations of $\pi$ and its gradients
- No MCMC or importance sampling
- In general non-convex, unless $\pi$ is log-concave
- Key steps: (1) parameterize, (2) optimize


## Low-dimensional structure of transport maps

Underlying challenge: maps in high dimensions

- Essential trade-off between expressiveness and computational effort/tractability!


## Low rank structure

(See [BBZSM 2020] for details.)

- Let $U=\left[U_{r} U_{\perp}\right] \in \mathbb{R}^{d \times d}$ be a unitary matrix, with $U_{r} \in \mathbb{R}^{d \times r}$. A lazy map $T: \mathbb{R}^{d} \rightarrow \mathbb{R}^{d}$ takes the form:

$$
T(z)=U_{r} \tau\left(z_{1}, \ldots, z_{r}\right)+U_{\perp} z_{\perp}
$$

for some diffeomorphism $\tau: \mathbb{R}^{r} \rightarrow \mathbb{R}^{r}$.

- Map $T \in \mathcal{T}_{r}(U)$ departs from the identity only on an $r$-dimensional subspace
- Proposition: For any lazy map $T \in \mathcal{T}_{r}(U)$, there exists a strictly positive function $f: \mathbb{R}^{r} \rightarrow \mathbb{R}_{+}$such that

$$
T_{\sharp} \rho(x)=f\left(U_{r}^{\top} x\right) \rho(x),
$$

for all $x \in \mathbb{R}^{d}$ where $\rho=\mathcal{N}\left(0, \mathbf{I}_{d}\right)$. Conversely, any density of the form $f\left(U_{r}^{\top} x\right) \rho(x)$ for some $f: \mathbb{R}^{r} \rightarrow \mathbb{R}_{+}$admits a lazy map representation.

## Discovering structure in $\pi$ before optimization

How to find a good $U_{r}$ ?

- Define

$$
H_{\pi}:=\mathbb{E}_{\pi}\left[\left(\nabla \log \frac{\pi}{\rho}\right)\left(\nabla \log \frac{\pi}{\rho}\right)^{\top}\right]
$$

- Let $\left(\lambda_{i}, u_{i}\right)$ be the $i$ th eigenpair of $H_{\pi}$ and put $U_{r}=\left[u_{1} u_{2} \cdots u_{r}\right]$.
- From previous results: There exists a map $T^{\star} \in \mathcal{T}_{r}(U)$ such that

$$
\mathcal{D}_{K L}\left(\pi \| T_{\sharp}^{\star} \rho\right) \leq \frac{1}{2}\left(\lambda_{r+1}+\ldots+\lambda_{d}\right) .
$$

## Discovering structure in $\pi$ before optimization

How to find a good $U_{r}$ ?

- Define

$$
H_{\pi}:=\mathbb{E}_{\pi}\left[\left(\nabla \log \frac{\pi}{\rho}\right)\left(\nabla \log \frac{\pi}{\rho}\right)^{\top}\right]
$$

- Let $\left(\lambda_{i}, u_{i}\right)$ be the $i$ th eigenpair of $H_{\pi}$ and put $U_{r}=\left[u_{1} u_{2} \cdots u_{r}\right]$.
- From previous results: There exists a map $T^{\star} \in \mathcal{T}_{r}(U)$ such that

$$
\mathcal{D}_{K L}\left(\pi \| T_{\sharp}^{\star} \rho\right) \leq \frac{1}{2}\left(\lambda_{r+1}+\ldots+\lambda_{d}\right) .
$$

- Good approximation when the spectrum of $H_{\pi}$ decays quickly
- $T^{\star}$ uses a ridge approximation of the likelihood $\frac{\mathrm{d} \pi}{\mathrm{d} \rho} \approx f^{\star}\left(U_{r}^{\top} x\right)$, with optimal profile function $f^{\star}\left(z_{r}\right)=\mathbb{E}_{X \sim \rho}\left[\left.\frac{\pi(X)}{\rho(X)} \right\rvert\, U_{r}^{\top} X=z_{r}\right]$.


## Error bound after optimization ("trace diagnostic")

Consider the matrix

$$
H_{T \sharp}:=\mathbb{E}_{T \sharp}\left[\left(\nabla \log \frac{T^{\sharp} \pi}{\rho}\right)\left(\nabla \log \frac{T^{\sharp} \pi}{\rho}\right)^{\top}\right]
$$

Then

$$
\mathcal{D}_{K L}\left(\pi \| T_{\sharp} \rho\right) \leq \frac{1}{2} \operatorname{Tr}\left(H_{T}\right) .
$$

Limiting case: if $T^{\sharp} \pi=\rho$, then $H_{T}=\mathbf{0}$ and $\mathcal{D}_{K L}\left(\pi \| T_{\sharp} \rho\right)=0$.

## Bound on the forward KL divergence for a given map.

## Layers of lazy maps

- What if $\left(\lambda_{i}\right)$ do not decay quickly? What if we are limited to small $r$ ?
- Answer: build a composition of lazy maps, via a greedy construction

$$
\mathfrak{T}_{\ell}=T_{1} \circ T_{2} \circ \cdots \circ T_{\ell}
$$

- Algorithm ("deeply lazy" maps):
- Given $\left(\pi, \rho, r_{1}\right)$ : compute $H_{\pi}$ and construct a first lazy map $T_{1}$
- Pull back $\pi$ by $T_{1}: \pi_{2}:=\left(T_{1}^{-1}\right)_{\sharp} \pi$
- Given $\left(\pi_{2}, \rho, r_{2}\right)$ : compute $H_{\pi_{2}}$ and construct a next lazy map $T_{2} \ldots$
- Generic iteration: at stage $\ell$, build a lazy map to the pullback $\pi_{\ell}:=\left(T_{1} \circ T_{2} \circ \cdots \circ T_{\ell-1}\right)_{\sharp}^{-1} \pi$
- Stop when $\frac{1}{2} \operatorname{Tr}\left(H_{\pi_{\ell}}\right)<\epsilon$


## Layers of lazy maps

Example: rotated "banana" target distribution, $r=1$ maps


## Example: log-Gaussian Cox process



Field $\boldsymbol{\Lambda}^{\star}$ and observations $\mathbf{y}^{\star}$


Realizations of $\boldsymbol{\Lambda} \sim \boldsymbol{\pi}_{\boldsymbol{\Lambda} \mid \mathrm{y}^{\star}}$

## Example: log-Gaussian Cox process

- Parameter dimension $n=4096,30$ observations; fixed ranks $r$



## Example: elliptic PDE Bayesian inverse problem

$$
\left\{\begin{array}{l}
\nabla \cdot\left(e^{\kappa(x)} \nabla u(\mathbf{x})\right)=0, \quad \text { for } \mathbf{x} \in \mathcal{D}:=[0,1]^{2}, \\
u(\mathbf{x})=0 \text { for } x_{1}=0, u(\mathbf{x})=1 \text { for } x_{1}=1, \frac{\partial u(\mathbf{x})}{\partial \mathbf{n}}=0 \text { for } x_{2} \in\{0,1\}
\end{array}\right.
$$

- Infer $\kappa(\mathbf{x})$, discretized with $n=2601$ parameters; 81 observations; lazy maps of $r \leq 4$ and polynomial degree up to 2

$u(\mathbf{x})$ and observations


Convergence


Posterior realizations of $\kappa(\mathbf{x})$

## Summary

- Identify and exploit low-dimensional structure in "updates" between distributions (from prior to posterior, from reference to target):
- Derive an upper bound on the forward KL divergence
- Minimize this upper bound using PCA on $\nabla \log \mathcal{L}_{y}$
- Better performance than heuristic gradient-based methods (e.g., likelihood-informed subspace or active subspaces)
- Transport methods: exploiting the pullback distribution
- Compositions of low-dimensional maps, constructed greedily ("deeply lazy" maps)


## Summary

- Identify and exploit low-dimensional structure in "updates" between distributions (from prior to posterior, from reference to target):
- Derive an upper bound on the forward KL divergence
- Minimize this upper bound using PCA on $\nabla \log \mathcal{L}_{y}$
- Better performance than heuristic gradient-based methods (e.g., likelihood-informed subspace or active subspaces)
- Transport methods: exploiting the pullback distribution
- Compositions of low-dimensional maps, constructed greedily ("deeply lazy" maps)


## Thanks for your attention!

## References

- M. Brennan, D. Bigoni, O. Zahm, A. Spantini, Y. Marzouk. "Greedy inference with structure-exploiting lazy maps." NeurIPS 2020.
- O. Zahm, T. Cui, K. Law, A. Spantini, Y. Marzouk. "Certified dimension reduction in nonlinear Bayesian inverse problems." arXiv:1807.03712v3, 2021.
- T. Cui, O. Zahm, "Data-free likelihood-informed dimension reduction of Bayesian inverse problems." Inverse Problems, 2021.
- T. Cui, X. Tong, "A unified performance analysis of likelihood-informed subspace methods." arXiv:2101.02417, 2021.
- O. Zahm, P. Constantine, C. Prieur, Y. Marzouk. "Gradient-based dimension reduction of multivariate vector-valued functions," SISC, 2020.
- A. Spantini, D. Bigoni, Y. Marzouk. "Inference via low-dimensional couplings." JMLR 19(66): 1-71, 2018.
- P. Constantine, C. Kent, T. Bui-Thanh. "Accelerating Markov chain Monte Carlo with active subspaces." SISC, 2016.
- A. Spantini, A. Solonen, T. Cui, J. Martin, L. Tenorio, Y. Marzouk, "Optimal low-rank approximations of Bayesian linear inverse problems," SISC, 2015.
- T. Cui, J. Martin, Y. Marzouk, A. Solonen. A. Spantini, "Likelihood-informed dimension reduction for nonlinear inverse problems," Inverse Problems, 2014.


## Approximation of $\pi_{\text {pos }}^{*}(x) \propto \quad \pi_{\mathrm{pr}}(x)$

- The conditional expectation $\mathbb{E}_{\mathrm{pr}}\left(\mathcal{L}_{y} \mid P_{r} x\right)$ can be expressed as

$$
x \mapsto \int \mathcal{L}_{y}\left(P_{r} x+\left(I_{d}-P_{r}\right) z\right) \pi_{\mathrm{pr}}\left(z \mid P_{r} x\right) \mathrm{d} z
$$

where $\pi_{\mathrm{pr}}\left(\cdot \mid P_{r} x\right)$ denotes the conditional prior, which depends on $x$.

## Approximation of $\pi_{\text {pos }}^{*}(x) \propto$ $\pi_{\mathrm{pr}}(x)$

- The conditional expectation $\mathbb{E}_{\mathrm{pr}}\left(\mathcal{L}_{y} \mid P_{r} x\right)$ can be expressed as

$$
x \mapsto \int \mathcal{L}_{y}\left(P_{r} x+\left(I_{d}-P_{r}\right) z\right) \pi_{\mathrm{pr}}\left(z \mid P_{r} x\right) \mathrm{d} z
$$

where $\pi_{\mathrm{pr}}\left(\cdot \mid P_{r} x\right)$ denotes the conditional prior, which depends on $x$.

- Consider the following Monte Carlo estimate

$$
\widetilde{\mathcal{L}}: x \mapsto \frac{1}{M} \sum_{i=1}^{M} \mathcal{L}_{y}\left(P_{r} x+\left(I_{d}-P_{r}\right) Z_{i}\right) \quad, \quad Z_{i} \stackrel{\text { iid }}{\sim} \pi_{\mathrm{pr}}
$$

In general, $\widetilde{\mathcal{L}}\left(P_{r} x\right)$ is a biased estimator for $\mathbb{E}_{\mathrm{pr}}\left(\mathcal{L}_{y} \mid P_{r} x\right)$.

## Approximation of $\pi_{\text {pos }}^{*}(x) \propto$

 $\pi_{\mathrm{pr}}(x)$- The conditional expectation $\mathbb{E}_{\mathrm{pr}}\left(\mathcal{L}_{y} \mid P_{r} x\right)$ can be expressed as

$$
x \mapsto \int \mathcal{L}_{y}\left(P_{r} x+\left(I_{d}-P_{r}\right) z\right) \pi_{\mathrm{pr}}\left(z \mid P_{r} x\right) \mathrm{d} z
$$

where $\pi_{\mathrm{pr}}\left(\cdot \mid P_{r} x\right)$ denotes the conditional prior, which depends on $x$.

- Consider the following Monte Carlo estimate

$$
\widetilde{\mathcal{L}}: x \mapsto \frac{1}{M} \sum_{i=1}^{M} \mathcal{L}_{y}\left(P_{r} x+\left(I_{d}-P_{r}\right) Z_{i}\right) \quad, \quad Z_{i} \stackrel{\mathrm{iid}}{\sim} \pi_{\mathrm{pr}}
$$

In general, $\widetilde{\mathcal{L}}\left(P_{r} x\right)$ is a biased estimator for $\mathbb{E}_{\mathrm{pr}}\left(\mathcal{L}_{y} \mid P_{r} x\right)$.

## Proposition

The random distribution $\widetilde{\pi}_{\mathrm{pos}}(x) \propto \widetilde{\mathcal{L}}\left(P_{r} x\right) \pi_{\mathrm{pr}}(x)$ is such that

$$
\mathbb{E}\left(D_{\mathrm{KL}}\left(\pi_{\mathrm{pos}}^{*} \| \widetilde{\pi}_{\mathrm{pos}}\right)\right) \lesssim\left(C_{1}+\frac{C_{2}}{M}\right) \mathcal{R}_{\pi_{\mathrm{pos}}}\left(P_{r}\right)
$$

## Convergence of the greedy construction

## Theorem (BBZSM21)

Let $U^{1}, U^{2}, \ldots$ be a sequence of unitary matrices. For any $\ell \geq 1$, let $T_{\ell} \in \mathcal{T}_{r}\left(U^{\ell}\right)$ be a lazy map that minimizes $\mathcal{D}_{K L}\left(\pi_{\ell-1} \|\left(T_{\ell}\right)_{\sharp} \rho\right)$, where $\pi_{\ell-1}=\left(T_{1} \circ \ldots \circ T_{\ell-1}\right)^{\sharp} \pi$. If there exists $0<t \leq 1$ such that for any $\ell \geq 1$

$$
\mathcal{D}_{K L}\left(\left(U_{r}^{\ell \top}\right)_{\sharp} \pi_{\ell-1} \| \rho_{r}\right) \geq t \sup _{\substack{U \in \mathbb{R}^{d \times d} \\ \text { s.t. } U U^{\top}=I_{d}}} \mathcal{D}_{K L}\left(\left(U_{r}^{\top}\right)_{\sharp} \pi_{\ell-1} \| \rho_{r}\right),
$$

then $\left(T_{1} \circ \ldots \circ T_{\ell}\right)_{\sharp} \rho$ converges weakly to $\pi$.

## Theorem (BBZSM21)

Let $U^{1}, U^{2}, \ldots$ be a sequence of unitary matrices. For any $\ell \geq 1$, let $T_{\ell} \in \mathcal{T}_{r}\left(U^{\ell}\right)$ be a lazy map that minimizes $\mathcal{D}_{K L}\left(\pi_{\ell-1} \|\left(T_{\ell}\right)_{\sharp} \rho\right)$, where $\pi_{\ell-1}=\left(T_{1} \circ \ldots \circ T_{\ell-1}\right)^{\sharp} \pi$. If there exists $0<t \leq 1$ such that for any $\ell \geq 1$

$$
\mathcal{D}_{K L}\left(\left(U_{r}^{\ell \top}\right)_{\sharp} \pi_{\ell-1} \| \rho_{r}\right) \geq t \sup _{\substack{U \in \mathbb{R}^{d \times d} \\ \text { s.t. } U U^{\top}=I_{d}}} \mathcal{D}_{K L}\left(\left(U_{r}^{\top}\right)_{\sharp} \pi_{\ell-1} \| \rho_{r}\right),
$$

then $\left(T_{1} \circ \ldots \circ T_{\ell}\right)_{\sharp} \rho$ converges weakly to $\pi$.

## Comments:

- This is a sufficient, not necessary, condition for convergence
- $t=1$ corresponds to an "ideal" greedy algorithm, but suboptimal choices for $U^{\ell}$ corresponding to $0<t<1$ are also sufficient
- Bound should apply simultaneously to all layers

