

# Chernoff's distribution and differential equations of parabolic and Airy type

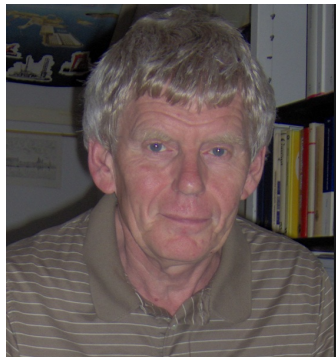
Piet Groeneboom, Delft University

joint work with:

Steve Lalley, University of Chicago,  
Nico Temme, CWI, Amsterdam

Banff, January-February, 2018

## Steven Lalley and Nico Temme



Analytical characterization of distribution of  
 $Z = \operatorname{argmax}\{W(t) - t^2 : t \in \mathbb{R}\}$ ?



Figure: Herman Chernoff

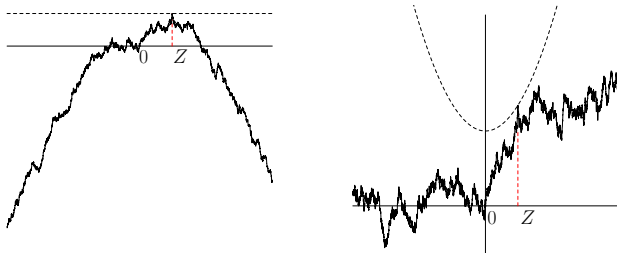


Figure: Left:  $Z$  and  $t \mapsto W(t) - t^2$ , right:  $Z$  and  $t \mapsto W(t)$ .

## Chernoff's heat equation (Chernoff (1964))

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If  $x < s^2$ :

$$\begin{aligned} u(s, x) = & \mathbb{P} \{ W(t) > t^2 \text{ for some } t > s + \epsilon \mid W(s) = x \} + \\ & + \mathbb{P} \{ W(t) > t^2 \text{ for some } t \in (s, s + \epsilon], \\ & \quad W(t) \leq t^2, \forall t > s + \epsilon \mid W(s) = x \} \end{aligned}$$

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## Distribution location of maximum (Chernoff (1964))

- Let  $M_h = \max_{t \in [s-h, s]} W(t)$ . Then:

$$\max_{t \in [s-h, s]} \{W(t) - t^2\} = M_h - s^2 + O(h)$$

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Using the spatial homogeneity of Brownian motion we get, taking  $\widetilde{W}(t) = W(t) - M_h + s$ ,

$$\begin{aligned} & \mathbb{P} \left\{ \max_{t \geq s} \{W(t) - t^2\} > M_h - s^2 \mid W(s), W(s-h), M_h \right\} \\ &= \mathbb{P} \left\{ \max_{t \geq s} \{\widetilde{W}(t) - t^2\} > 0 \mid \widetilde{W}(s) = W(s) - M_h + s^2 \right\} \end{aligned}$$

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## Distribution location of maximum (Chernoff (1964))

- Similarly:

$$\begin{aligned} & \mathbb{P} \left\{ \max_{t \leq s-h} \{W(t) - t^2\} \geq M_h - (s-h)^2 \mid W(s-h), M_h \right\} \\ &= u(-s, s^2) \quad (= 1) - \{M_h - W(s-h)\} \partial_2 u(-s, s^2) + O_p(h). \end{aligned}$$



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## Chernoff's theorem

Theorem (Chernoff (1964))

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Then the density  $f_Z$  of  $Z = \operatorname{argmax}\{W(x) - x^2\}$  is given by:

$$f_Z(s) = \frac{1}{2} \partial_2 u(-s, s^2) \partial_2 u(s, s^2).$$

where  $u(s, x)$  solves the heat equation:

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subject to:

$$u(s, x) = 1, \quad x \geq s^2, \quad u(s, x) \rightarrow 0, \quad x \rightarrow -\infty.$$

## Computation of density

Original computations of this density were based on numerically solving Chernoff's heat equation.

But (Groeneboom (1984)):

$$\partial_2 u(-s, s^2) \sim c_1 \exp \left\{ -\frac{2}{3}s^3 - cs \right\}, \quad s \rightarrow \infty,$$

where  $c \approx 2.9458$  and  $c_1 \approx 2.2638$ . This entails that a numerical solution of this partial differential equation on a grid will not give a really accurate solution!

## Analytic approach



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Let  $Q^{(s,x)}$  is the probability measure of  $t \mapsto W(t) - t^2$ , starting at  $x < 0$  at time  $s$   
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Chernoff's result: the density of  $\operatorname{argmax}_t \{W(t) - t^2\}$  is given by:

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where

$$\begin{aligned} u_2(s) &= \lim_{x \uparrow 0} \frac{\partial}{\partial x} u(s, x) = \lim_{x \uparrow 0} \frac{\partial}{\partial x} Q^{(s,x)} \{X_t \geq 0, \text{ for some } t \geq s\} \\ &= - \lim_{x \uparrow 0} \frac{\partial}{\partial x} Q^{(s,x)} \{X_t < 0, \forall t \geq s\} \\ &= - \lim_{x \uparrow 0} \frac{\partial}{\partial x} Q^{(s,x)} \{\tau_0 = \infty\}, \end{aligned}$$

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## Cameron-Martin-Girsanov

- (1) under  $P^{(s,x)}$ ,  $\{X_t = W_t : t \geq s\}$  is standard Brownian motion with  $X_s = x$ ,
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Hence, if  $x < 0$ ,  $s < t$ , and  $\tau_0$  is the first time  $X_t$  hits zero:

$$\begin{aligned} & Q^{(s,x)} \{ \tau_0 \in dt \} \\ &= \exp \left\{ 2sx - \frac{2}{3}(t^3 - s^3) \right\} \\ & \cdot E^{P^{(s,x)}} \left\{ \exp \left\{ 2 \int_s^t X_u du \right\} \mid \tau_0 = t \right\} P^{(s,x)} \{ \tau_0 \in dt \}. \end{aligned}$$

## Airy functions

Let, for  $\lambda > 0$ ,  $u_\lambda$  be the unique non-negative solution of the boundary problem

$$\frac{1}{2}u''(x) - (\lambda - 2x)u(x) = 0, \quad x < 0, \quad \lim_{x \uparrow 0} u(x) = 1, \quad u(x) \leq 1, \quad x \leq 0.$$

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Consider the process

$$Y_t = e^{-\int_{v=s}^t (\lambda - 2X_v) dv} u_\lambda(X_t),$$

where  $X_t$  is standard Brownian motion, starting at  $x < 0$  at time  $s$ .

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$Y_t$  is a local martingale (**Ito's formula**). Hence, if  $t > s$ .

$$E^{P(s,x)} e^{-\int_{v=s}^{\tau_0} (\lambda - 2X_v) dv} = E^{P(s,x)} Y_{\tau_0} = Y_s = u_\lambda(x),$$

# Airy functions

Conclusion:

$$\int_{t \in (s, \infty)} e^{-\lambda(t-s)} E^{P(s,x)} \left\{ e^{\int_{v=s}^t 2X_v dv} \mid \tau_0 = t \right\} P^{(s,x)} \{ \tau_0 \in dt \}$$
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So we can compute:

$$Q(s,x) \{ \tau_0 \in dt \}$$
$$= \exp \left\{ 2sx - \frac{2}{3}(t^3 - s^3) \right\}$$
$$\cdot E^{P(s,x)} \left\{ e^{\int_{v=s}^t 2X_v dv} \mid \tau_0 = t \right\} P(s,x) \{ \tau_0 \in dt \},$$

where  $Q(s,x)$  is the probability measure of  $\{W_t - t^2 : t \geq s\}$ , starting at  $x$  at time  $s$ .

## Inverse Laplace transforms

Inversion of the Laplace transform along the imaginary axis:

$$\begin{aligned} Q^{(s,x)} \{ \tau_0 < \infty \} &= \int_{t \in (s, \infty)} Q^{(s,x)} \{ \tau_0 \in dt \} \\ &= \frac{e^{2sx + \frac{2}{3}s^3}}{2\pi} \int_{v=-\infty}^{\infty} \frac{\text{Ai}(2^{-1/3}iv - 4^{1/3}x)}{\text{Ai}(2^{-1/3}iv)} \int_{t=0}^{\infty} e^{itv - \frac{2}{3}(s+t)^3} dt dv. \end{aligned}$$

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So we would be done if we can deal with the properties of the integral in the last line, since Chernoff's function  $u(s, x)$  satisfies

$$u(s, x) = Q^{(s,x)} \{ \tau_0 < \infty \}.$$

Taking the special case  $s = 0$ , we get:

$$\frac{1}{\pi} \int_{t=0}^{\infty} e^{itv - \frac{1}{3}t^3} dt = \text{Hi}(iv),$$

where **Hi** denotes **Scorer's function Hi**.

## Difficulties of direct approach

Scorer's function  $\text{Hi}$  has the asymptotic expansion, as  $|z| \rightarrow \infty$ ,

$$\text{Hi}(z) \sim -\frac{1}{\pi z} \sum_{k=0}^{\infty} \frac{(3k)!}{k!(3z^3)^k}, \quad |\text{ph}(-z)| < \frac{2}{3}\pi - \delta$$

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This means that

$$\int_{v=-\infty}^{\infty} \left\{ \frac{\text{Ai}(2^{-1/3}iv - 4^{1/3}x)}{\text{Ai}(2^{-1/3}iv)} \int_{t=0}^{\infty} e^{itv - \frac{2}{3}(s+t)^3} dt \right\} dv$$

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has a non-integrable integrand if  $x = 0$ . Whereas in fact:

$$\begin{aligned} & \lim_{x \uparrow 0} \int_{v=-\infty}^{\infty} \frac{\text{Ai}(2^{-1/3}iv - 4^{1/3}x)}{\text{Ai}(2^{-1/3}iv)} \int_{t=0}^{\infty} e^{itv - \frac{2}{3}(s+t)^3} dt dv \\ &= \lim_{x \uparrow 0} Q^{(s,x)} \{ \tau_0 < \infty \} = 1. \end{aligned}$$

## Indirect approach

For this reason the limit

$$Q^{(s,x)} \{ \tau_0 = \infty \} = \lim_{t \rightarrow \infty} Q^{(s,x)} \{ \tau_0 > t \}$$

was computed by first determining the transition density

$$Q^{(s,x)} \left\{ X_t^\partial \in dy \right\}, \quad t > s, \quad x, y < 0,$$

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$$\begin{aligned} & Q^{(s,x)} \{ \tau_0 = \infty \} \\ &= c_{s,x} \int_{v=-\infty}^{\infty} e^{-isv} \frac{\text{Ai}(i\xi) \text{Bi}(i\xi - 4^{1/3}x) - \text{Ai}(i\xi - 4^{1/3}x) \text{Bi}(i\xi)}{\text{Ai}(i\xi)} dv, \end{aligned}$$

where  $\xi = 2^{-1/3}v$  and  $c_{s,x} = 4^{-1/3} \exp\{(2/3)s^3 + 2sx\}$ .

Since we also know:

$$\begin{aligned}
 & Q^{(s,x)} \{ \tau_0 = \infty \} \\
 &= 1 - \frac{e^{2sx + \frac{2}{3}s^3}}{2\pi} \int_{v=-\infty}^{\infty} \frac{\text{Ai}(2^{-1/3}iv - 4^{1/3}x)}{\text{Ai}(2^{-1/3}iv)} \int_{t=0}^{\infty} e^{itv - \frac{2}{3}(s+t)^3} dt dv,
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If we could prove this relation analytically, we wouldn't need to introduce the process  $X_t^\partial$ , and we would not need the complicated computations in the appendix of Groeneboom (1989).

### Lemma (PDE for first expression)

Let the function  $f : \mathbb{R} \times (-\infty, 0) \rightarrow \mathbb{R}$  be defined by

$$f(s, x) = \frac{1}{2\pi} \int_{v=-\infty}^{\infty} \frac{\text{Ai}(i\xi - 4^{1/3}x)}{\text{Ai}(i\xi)} \int_{t=0}^{\infty} e^{itv - \frac{2}{3}(s+t)^3} dt dv,$$

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Moreover  $0 \leq f(s, x) \leq e^{-2sx - \frac{2}{3}s^3}$ ,  $\lim_{s \rightarrow \infty} f(s, x) = 0$ , and

$$\lim_{x \uparrow 0} f(s, x) = e^{-\frac{2}{3}s^3}, \quad \lim_{x \rightarrow -\infty} e^{2sx} f(s, x) = 0, \quad s \in \mathbb{R}.$$

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Moreover,

$$\lim_{x \uparrow 0} g(s, x) = 0, \quad \lim_{x \rightarrow -\infty} e^{2sx} g(s, x) = e^{-\frac{2}{3}s^3}, \quad s > 0.$$

## Theorem

(i) Let  $f$  and  $g$  be as in the preceding lemmas. Then:

$$f(s, x) = e^{-2sx - \frac{2}{3}s^3} - g(s, x), \quad s \in \mathbb{R}, \quad x \leq 0.$$

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$$Q^{(s,x)}\{\tau_0 < \infty\} = e^{2sx + \frac{2}{3}s^3} f(s, x),$$

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Proof.

$$f(s, x) + g(s, x) - e^{-2sx - \frac{2}{3}s^3} = 0, \quad s \in \mathbb{R}, \quad x \leq 0.$$



Proof of  $h(s, x) = f(s, x) + g(s, x) - e^{-2sx} - \frac{2}{3}s^3 \equiv 0$ .

One can show  $h(s, 0) = 0$ ,  $s \in \mathbb{R}$ , and:

$$\lim_{x \rightarrow -\infty} h(s, x) = 0, \forall s \in \mathbb{R}, \quad \text{and} \quad \lim_{s \rightarrow \infty} h(s, x) = 0, \forall x < 0.$$



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Consider an infinite rectangle  $R_c = \{(s, x) : s \geq c, x \leq 0\}$ ,  $c \in \mathbb{R}$ . Suppose  $h$  attains a strictly positive maximum over  $R_c$  at an interior point  $(s_0, x_0) \in R_c^0$ . Then  $\partial_1 h(s_0, x_0) = 0$ . Hence

$$0 = \partial_1 h(s_0, x_0) = -\frac{1}{2}\partial_2^2 h(s_0, x_0) - 2x_0 h(s_0, x_0).$$

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This implies:  $\partial_2^2 h(s_0, x_0) = -4x_0 h(s_0, x_0) > 0$ , since  $x_0 < 0$  and  $h(s_0, x_0) > 0$ .

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Proof of  $h(s, x) = f(s, x) + g(s, x) - e^{-2sx} - \frac{2}{3}s^3 \equiv 0$ .

So a strictly positive maximum or strictly negative minimum over  $R_c$  can only be attained on the line  $s = c$ .

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Suppose that a strictly positive maximum is attained at the point  $(c, x_0)$ , where  $x_0 < 0$ . Then we must have:  $\partial_1 h(c, x_0) \leq 0$ , implying by the partial differential equation for  $h$ :

$$\partial_2^2 h(c, x_0) \geq -4x_0 h(c, x_0) > 0,$$

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**Conclusion:**  $h$  is identically zero on  $R_c$ . Since the argument holds for all  $c \in \mathbb{R}$ , we get that the function  $h$  is identically zero on  $\mathbb{R} \times (-\infty, 0]$ .

Theorem (Groeneboom (1984), Daniels and Skyrme (1985))

The probability density  $f$  of the location of the maximum of the process  $t \mapsto W(t) - t^2$ ,  $t \in \mathbb{R}$ , is given by

$$f(s) = \frac{1}{2}g(s)g(-s),$$

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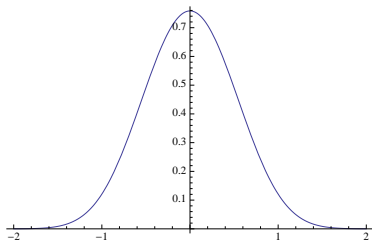
Distribution of the **maximum itself**:

Janson, Louchard, and Martin-Löf (2010), Groeneboom (2010)  
Groeneboom and Temme (2011) and [Groeneboom, Lalley, and Temme \(2013\)](#) (joint density of max and argmax).

## Henry Daniels



Density of  $Z = \operatorname{argmax}\{W(t) - t^2, t \in \mathbb{R}\}$



**Figure:** The density  $f_Z$  of the location of the maximum  $Z$  of  $W(t) - t^2$ ,  $t \in \mathbb{R}$ .

Density of  $Z = \operatorname{argmax}\{W(t) - t^2, t \in \mathbb{R}\}$

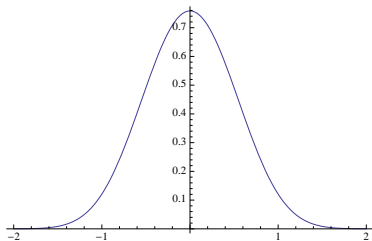


Figure: The density  $f_Z$  of the location of the maximum  $Z$  of  $W(t) - t^2, t \in \mathbb{R}$ .

Also:

$$\operatorname{var}(Z) = \frac{1}{3} \mathbb{E} \max_t \{W(t) - t^2\},$$

as proved in Groeneboom (2011) and Janson (2013), and (not using the relation with Airy functions) in Pimentel (2014).

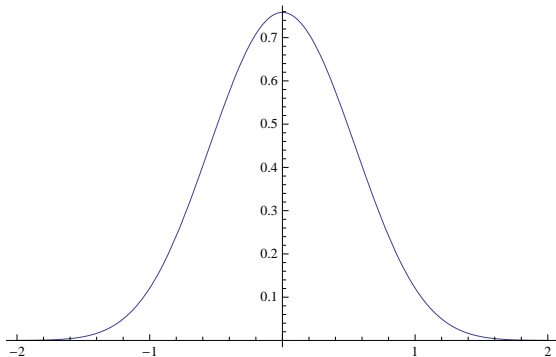
The density can be computed by two lines in Mathematica:

```
In[1]:= f[x_] := (1 / (2 * Pi)) * 2^(1 / 3) *  
         Re[NIntegrate[Exp[-I * u * x] / AiryAi[I * 2^(-1 / 3) * u], {u, -10, 10}]]
```

```
In[2]:= g[x_] := (1 / 2) * f[x] * f[-x]
```

```
In[3]:= Plot[g[x], {x, -2, 2}]
```

Out[3]=



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