

# The largest eigenvalue of finite rank deformation of Wigner matrices

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## 1 The model

$$M_N = X_N + A_N := \frac{1}{\sqrt{N}}W_N + A_N$$

where  $W_N$  is a  $N \times N$  Hermitian matrix such that  $(W_N)_{ii}$ ,  $\sqrt{2}\operatorname{Re}((W_N)_{ij})_{i < j}$ ,  $\sqrt{2}\operatorname{Im}((W_N)_{ij})_{i < j}$  are iid with common distribution  $\mu$ .  $\mu$  is assumed to be symmetric with variance  $\sigma^2$  and it satisfies a Poincaré inequality.

$A_N$  is a deterministic, Hermitian matrix.

**Example:**  $\mu = N(0; \sigma^2)$ ,  $X_N \sim GUE(N, \frac{\sigma^2}{N})$ .

## 2 Some known result in the non deformed case ( $A_N = 0$ )

- Convergence of the spectral measure  $\mu_{X_n} := \frac{1}{N} \sum_i \delta_{\lambda_i(X_N)}$  to the Wigner distribution  $\mu_{sc} = \frac{1}{2\pi\sigma^2} \sqrt{4\sigma^2 - x^2} 1_{[-2\sigma, 2\sigma]}$ .

- Convergence a.s. of  $\lambda_{max}(X_N)$  to  $2\sigma$  (the right endpoint of the support of the limiting distribution)
- Fluctuations (Tracy-Widom, Soshnikov)

$$\sigma^{-1}N^{2/3}(\lambda_{max}(X_N) - 2\sigma) \xrightarrow{\mathcal{L}} \text{T-W distribution } F_2$$

where the distribution  $F_2$  can be expressed with the Fredholm determinant of an operator associated to the Airy kernel.

### 3 The deformation

$A_N$  Hermitian of finite rank  $r$  (independent of  $N$ ) with eigenvalues  $\theta_i$  of multiplicity  $k_i$ ;  $\theta_1 > \theta_2 > \dots > \theta_J$ .

Convergence of the spectral measure to the semicircular distribution  $\mu_{sc}$ .

What about the extremal eigenvalues?

## 1) The Gaussian case (Péché)

Ex:  $\theta_1$  with multiplicity 1.

$$1) \text{ si } 0 \leq \theta_1 < \sigma, \quad \sigma^{-1} N^{2/3} (\lambda_{\max}(M_N) - 2\sigma) \xrightarrow{\mathcal{L}} F_2$$

$$2) \text{ si } \theta_1 = \sigma, \quad \sigma^{-1} N^{2/3} (\lambda_{\max}(M_N) - 2\sigma) \xrightarrow{\mathcal{L}} F_3$$

$$3) \text{ si } \theta_1 > \sigma, \quad N^{1/2} (\lambda_{\max}(M_N) - \rho_{\theta_1}) \xrightarrow{\mathcal{L}} \mathcal{N}(0, \sigma_{\theta_1}^2)$$

with  $\rho_{\theta_1} = \theta_1 + \frac{\sigma^2}{\theta_1} > 2\sigma$ .

## 2) The non Gaussian case for a particular $A_N$ (Féral-Péché)

$A_N$  is the deformation defined by  $(A_N)_{ij} = \frac{\theta}{N}$ , so that  $r = 1$  and  $\theta_1 = \theta$ .

Same TCL as in the Gaussian case, universality of the fluctuations (independent of  $\mu$ , the distribution of the entries).

### 3) The non Gaussian case, $A_N$ general

**Theorem 1** *a.s. behaviour of the spectrum of  $M_N$ .  
Let  $J_{+\sigma}$  (resp.  $J_{-\sigma}$ ) be the number of  $j$ 's such that  
 $\theta_j > \sigma$  (resp.  $\theta_j < -\sigma$ ).*

(a)  $\forall 1 \leq j \leq J_{+\sigma}, \forall 1 \leq i \leq k_j,$

$$\lambda_{k_1+\dots+k_{j-1}+i}(M_N) \longrightarrow \rho_{\theta_j} \quad a.s.,$$

(b)  $\lambda_{k_1+\dots+k_{J_{+\sigma}}+1}(M_N) \longrightarrow 2\sigma \quad a.s.,$

(c)  $\lambda_{k_1+\dots+k_{J-J_{-\sigma}}}(M_N) \longrightarrow -2\sigma \quad a.s.,$

(d)  $\forall j \geq J - J_{-\sigma} + 1, \forall 1 \leq i \leq k_j,$

$$\lambda_{k_1+\dots+k_{j-1}+i}(M_N) \longrightarrow \rho_{\theta_j} \quad a.s.$$

**Remark:** Same result as in the sample covariance matrices (Bai-Silverstein, Baik-Silverstein)

## 4 Elements of Proof of Theorem 1

**Step 1** Prove that a.s.

$$\text{Spect}(M_N) \subset K_\sigma(\theta_1, \dots, \theta_J) + [-\epsilon, +\epsilon] \quad (1)$$

for  $N$  large, where  $K_\sigma(\theta_1, \dots, \theta_J) :=$

$$\left\{ \rho_{\theta_J}; \dots ; \rho_{\theta_{J-J_{-\sigma}+1}} \right\} \cup [-2\sigma; 2\sigma] \cup \left\{ \rho_{\theta_{J+\sigma}}; \dots ; \rho_{\theta_1} \right\}.$$

**Tool:** The Stieltjes transform: for  $z \in \mathbb{C} \setminus \mathbb{R}$ , define  $g_N(z) = \text{tr}_N(G_N(z))$  where  $G_N(z) = (zI_N - M_N)^{-1}$  is the resolvent of  $M_N$ . We set  $h_N(z) = \mathbb{E}[g_N(z)]$ .

$$g_N(z) = \int \frac{1}{z-x} d\mu_{M_N}(x); \quad h_\sigma(z) = \int \frac{1}{z-x} d\mu_{sc}(x).$$

**Aim:** Obtain a precise estimate

$$h_\sigma(z) - h_N(z) + \frac{1}{N}L_\sigma(z) = O\left(\frac{1}{N^2}\right) \quad (2)$$

where  $L_\sigma$  is the Stieltjes transform of a distribution  $\eta$  with compact support in  $K_\sigma$ .

With the help of the inverse Stieltjes transform,

$$\mathbb{E}[\text{tr}_N(\varphi(M_N))] = \int \varphi(x) d\mu_{sc}(x) + \frac{1}{N} \int \varphi(x) d\eta(x) + O\left(\frac{1}{N^2}\right),$$

for  $\varphi$  smooth with compact support;

and some variance estimates, we deduce from (2)

$$\text{tr}_N 1_{c_{K_\sigma^\epsilon}(\theta_1, \dots, \theta_J)}(M_N) = O(1/N^{\frac{4}{3}}) \text{ a.s.}$$

and therefore the inclusion of the spectrum (1).

**Proof of (2):**

**1) The Gaussian Case:**

- The Gaussian integration by parts formula:

$\phi : \mathbb{R} \rightarrow \mathbb{C}$ ,  $\xi$  *standard Gaussian*

$$\mathbb{E}(\xi\phi(\xi)) = \mathbb{E}(\phi'(\xi)).$$

$\Phi : \mathcal{H}_n(\mathbb{C}) \rightarrow \mathbb{C}$ ,  $H \in \mathcal{H}_n(\mathbb{C})$ ,

$$\frac{N}{\sigma^2} \mathbb{E}[\text{Tr}(X_N H) \Phi(X_N)] = \mathbb{E}[\Phi'(X_N) \cdot H]$$

Apply it for  $\Phi(X_N) = [(zI_N - X_N - A_N)^{-1}]_{kl} = G_N(z)_{kl}$  and  $H = E_{kl}$ ; then sum over  $k$  and  $l$ .

$$\rightarrow \sigma^2 \mathbb{E}[g_N^2(z)] - z \mathbb{E}[g_N(z)] + 1 + \frac{1}{N} \mathbb{E}[\text{Tr}(G_N(z) A_N)] = 0$$

$$\rightarrow \sigma^2 h_N^2(z) - z h_N(z) + 1 + \frac{1}{N} \mathbb{E}[\text{Tr}(G_N(z) A_N)] = O\left(\frac{1}{N^2}\right)$$

Recall that  $\sigma^2 h_\sigma^2(z) - z h_\sigma(z) + 1 = 0$ .

**Estimate for  $\mathbb{E}[\text{Tr}(G_N(z)A_N)]$ :**

$A_N = U^* \Lambda U$  where  $\Lambda$  is a diagonal matrix with entries  $\lambda_i \neq 0$  for  $i \leq r$ ,  $\lambda_i = 0$ ,  $i > r$ . We can show using

- The Gaussian integration by parts formula
- Some variance estimates
- $h_N(z) = h_\sigma(z) + O(\frac{1}{N})$

the estimate

$$\mathbb{E}[\text{Tr}(G_N(z)A_N)] = \sum_{i=1}^r \frac{\lambda_i}{z - \lambda_i - \sigma^2 h_\sigma(z)} + O(\frac{1}{N})$$

Set

$$R_G^{A_N}(z) = \sum_{i=1}^r \frac{\lambda_i}{z - \lambda_i - \sigma^2 h_\sigma(z)} = \sum_{\theta_i \neq 0} k_i \frac{\theta_i}{z - \theta_i - \sigma^2 h_\sigma(z)}.$$

Then,

$$\sigma^2 h_N^2(z) - z h_N(z) + 1 + \frac{1}{N} R_G^{A_N}(z) = O(\frac{1}{N^2})$$

leading to

$$h_N(z) - h_\sigma(z) + \frac{1}{N} L(z) = O(\frac{1}{N^2})$$

where  $L(z) = h_\sigma^{-1}(z) \mathbb{E}[(z - sc)^{-2}] R_G^{A_N}(z)$ .

### Question:

- $L$  Stieltjes transform of a distribution ?
- Support of this distribution?

$\longleftrightarrow$  Analyticity of  $L$  (+ conditions); set of singular points.

If  $|\theta_i| > \sigma$ ,  $x \in \mathbb{R} \setminus [-2\sigma, 2\sigma]$ ,

$$x - \theta_i - \sigma^2 h_\sigma(x) = 0 \iff x = \theta_i + \frac{\sigma^2}{\theta_i} := \rho_{\theta_i}.$$

## 2) The non Gaussian case

GIP replaced by: (Khorunzhy, Khoruzhenko, Pastur)

**Lemma 1** *Let  $\xi$  be a real-valued rv such that  $\mathbb{E}(|\xi|^{p+2}) < \infty$ . Let  $\phi : \mathbb{R} \rightarrow \mathbb{C}$  such that the first  $p+1$  derivatives are continuous and bounded. Then,*

$$\mathbb{E}(\xi \phi(\xi)) = \sum_{a=0}^p \frac{\kappa_{a+1}}{a!} \mathbb{E}(\phi^{(a)}(\xi)) + \epsilon$$

where  $\kappa_a$  are the cumulants of  $\xi$ ,  $|\epsilon| \leq C \sup_t |\phi^{(p+1)}(t)| \mathbb{E}(|\xi|^{p+2})$ .

Apply to  $\xi = \text{Re}((X_N)_{ij})$ ,  $\text{Im}((X_N)_{ij})$ ,  $(X_N)_{ii}$ , the odd cumulants vanish ( $\mu$  symmetric). One must consider the third derivative of  $\Phi = (G_N(z))_{kl}$ .



One obtains:

$$\sigma^2 h_N^2(z) - zh_N(z) + 1 + \frac{1}{N}R(z) = O\left(\frac{1}{N^2}\right)$$

where  $R(z) = R_G^{AN}(z) + \kappa_4 R_{\Phi'''}^0(z)$ .

$R_{\Phi'''}^0(z)$  is analytic on  $\mathbb{C} \setminus [-2\sigma, 2\sigma]$ .