# Tail bounds and extremal behavior of light-tailed perpetuities 

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## Perpetuities

By perpetuity we mean a random variable $R$ which satisfies the following distributional equation:

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R \stackrel{d}{=} M R+Q
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Assuming the first term is negligible and re-numbering $\left(Q_{n}, M_{n}\right)$

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And for the almost sure convergence to 0 of

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- This basic result has been re-proved and extended by a number of researchers, among others Goldie (1991), Grey (1994), Grincievičjus (1975) ...


## Light tails

In most applications, $P(\mid M>1)>0$ so that most of the time we are interested in the heavy-tails. However, the complementary case

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For example:

- in the context of record times of random random walks Vervaat (1972) studied the situation in which

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Q \equiv 1 \quad \text { and } \quad M \stackrel{d}{=} \operatorname{beta}(\alpha, 1) \stackrel{d}{=} U^{1 / \alpha}
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- Such perpetuities are nowadays called Vervaat perpetuities.


## Light tails, cont.

- in the special case $\alpha=1$ the density of Vervaat perpetuity is (up to normalizing constant) the Dickman function $\rho(u)$ appearing in number theory:

$$
\rho(u)=\lim _{n \rightarrow \infty} \frac{k_{n}(u)}{n}
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- other appearances of Dickman function are discussed in Hwang and Tsai (2001) and include the analysis of Quickselect algorithm, the degree of the largest irreducible factor in a random polynomial over finite field, and allele frequencies in some biological models.


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- Goldie and Grübel (1996) were the first to study light-tailed case in some generality (apparently, they were unaware of those earlier special results). They showed that:


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for some $\epsilon>0,0<c, C<\infty$ and for all $\delta \in(0, \epsilon]$ then

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\lim _{x \rightarrow \infty} \frac{\ln (P(R \geq x))}{x \ln x}=-\frac{1}{q} .
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- H. and Wesołowski (2009) extended these ideas to construct $M$ 's for which the corresponding $R$ satisfies, for example:

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Then, when $Q \equiv q>0$, for $c \in(0,1)$ and $x>q$ we have

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- Upper bound: Theorem (H. (2010)): There exist constants $c_{1}, c_{2}$ such that if $|Q| \leq q$ and $|M| \leq 1$ then for sufficiently large $x$ :

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- In particular, if $Q \equiv q>0$ and $0 \leq M \leq 1$ then

$$
\exp \left(\frac{2 \ln 2}{q} x \ln p_{q /(2 x)}\right) \leq P(R>x) \leq \exp \left(\frac{1}{4 q} x \ln p_{2 q / x}\right)
$$

## Comments on proof

- techniques for the cases $0 \leq M \leq 1$ and $P(M>1)>0$ are completely different.
- techniques previously used for an upper bound in the case $0 \leq M \leq 1$ were generally based on an iteration of the equation $R_{n} \stackrel{d}{=} M_{n} R_{n-1}+Q_{n}$ and they don't seem to work.
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- However, a proof of a lower bound of Goldie-Grübel may be used to yield an upper bound.
- Rough idea for the lower bound: for a small $\delta$, wait for the first time when $M_{k} \leq 1-\delta$. Up to that time bound the partial sums forming $R_{n}$ below by a geometric sum.
- For the upper bound: keep recording consecutive times when $M_{k} \leq 1-\delta$, bound above the partial sums by weighted sums of geometric r.v.'s and use exponential bounds for such sums (Goh, H. (2008)).


## Extremal behavior: heavy tails

We want to analyze the extremal behavior of $\left(R_{n}\right)$ i.e. look at the normalizing constant $a_{n}$ and $b_{n}$ so that

$$
a_{n}\left(\max _{0 \leq k \leq n} R_{n}-b_{n}\right)
$$

converges in distribution to a non-degenerate random variable. The theory for i.i.d. sequences $\left(R_{n}\right)$ is completely developed and goes back to Fisher-Tippett (1928) and Gnedenko (1943) and is presented e.g. in a classic Leadbetter, Lindren and Rootzén (1988). The situation is also well understood when $\left(R_{n}\right)$ is a stationary sequence. In our case, if $\left(R_{n}\right)$ converges in distribution to $R$ we can take $R_{0} \stackrel{d}{=} R$ and turn $\left(R_{n}\right)$ into a stationary sequence.

## Extremal behavior: heavy tails, cont.

de Haan, Resnick, Rootzén, de Vries (1989) showed that

- for $M, Q \geq 0$ under Kesten's conditions for the convergence of $\left(R_{n}\right)$ and $P(M>1)>0$ (which implies that $\left.P(R>x) \sim c x^{-\kappa}\right)$ we have

$$
\lim _{n \rightarrow \infty} P\left(\frac{R_{n}^{*}}{n^{1 / \kappa}} \leq x\right)=\exp \left(-c \theta x^{-1 / \kappa}\right)
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where $R_{n}^{*}=\max _{1 \leq k \leq n} R_{k}$.

- That is, there is a convergence to Type II (Frechét) distribution with normalizing constants $a_{n}=1 / n^{1 / \kappa}$ and $b_{n}=0$.


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- $\theta=\kappa \int_{1}^{\infty} P\left(\sup _{j \geq 1} \prod_{i=1}^{j} M_{i} \leq \frac{1}{y}\right) \frac{d y}{y^{k+1}}$, is the extremal index of the sequence $\left(R_{n}\right)$.
- the existence of such $\theta \in[0,1]$ (NOT assured in general even for stationary sequences) says that $R_{n}^{*}$ behaves like max of $\sim \theta n$ i.i.d. variables with the same marginal distribution.


## Extremal behavior: light tails

Theorem (H. (2010)): Let $R_{n}=M_{n} R_{n-1}+q$ where $q>0$, $0 \leq M \leq 1 M$ is non-degenerate, $P(M=0)=0$, and $\sup \{x: P(M>x)>0\}=1$. Then there exist $\left(a_{n}\right),\left(b_{n}\right)$ such that

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- the assumptions on $M$ are needed to exclude trivial cases and the case when $R$ (and hence each $R_{n}$ ) is geometric.
- We may take $b_{n}$ so that $b_{n} \ln p_{c / b_{n}}=-\Theta(\ln n)$ and

$$
a_{n}=\Theta\left(\frac{1}{b_{n} \ln p_{c / b_{n}}} f_{M}\left(1-\frac{c}{b_{n}}\right)-\ln p_{c / b_{n}}\right) \stackrel{*}{\sim}-\Theta\left(\ln p_{c / b_{n}}\right)
$$

where ' $\stackrel{*}{\sim}$ ' means 'often $\sim$ ' and $c$ is a constant.

- the extremal index (built-in in $\left.a_{n}, b_{n}\right)$ is $\theta=1-P(M=1)$.


## Open problems: tail behavior

- Get the asymptotics for $P(R>x)$ when $R \stackrel{d}{=} M R+q$, $0 \leq M \leq 1, q>0$.
Knowing the tail behavior would give the asymptotics of the normalizing constants $a_{n}, b_{n}$ in the limit theorem for the extremes.
- Get rid of the assumption $Q \equiv q$ and/or $|Q| \leq q$. Without that some of the basic cases when we know the tail behavior are not covered, e.g. the $\alpha$-stable distributions:

$$
R \stackrel{d}{=} 2^{-1 / \alpha}\left(R+R^{\prime}\right) \stackrel{d}{=} M R+Q ; M=2^{-1 / \alpha}, Q \stackrel{d}{=} 2^{-1 / \alpha} R
$$

or

$$
M \stackrel{d}{=} \beta\left(\alpha_{1}, \alpha_{2}\right), Q \stackrel{d}{=} \Gamma\left(\alpha_{2}, \gamma\right) \Longrightarrow R \stackrel{d}{=} \Gamma\left(\alpha_{1}+\alpha_{2}, \gamma\right) .
$$

## Thank you :)



Analysis and Probability, June 10-16, 2012 conference website: http://www.mimuw.edu.pl/~probanal

