# Sequential (Interactive) Testing in High Dimensions 

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BIRS 2012 - Interactive Information Theory Joint work with Rob Nowak

## Problem Setup

- Consider a support set $\mathcal{S} \subset\{1, \ldots, n\}$ with $|\mathcal{S}|=s \ll n$ and the random variables

$$
Y_{i} \sim\left\{\begin{array}{ll}
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$2 / 17$

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- Definition: take on average $m$ samples of each index:

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- This talk: find relationship between $(n, s, m)$ such that $\mathbb{P}(\hat{\mathcal{S}}=\mathcal{S}) \rightarrow 1$.



## Related work

A. Wald and J. Wolfowitz

Optimum character of the sequential probability ratio test. 1948. 1-dimensional simple binary hypothesis test

## E. Posner

Optimal Search Procedures. 1963. AWGN, $s=1$, SPRT
J. Haupt, R. Castro, and R. Nowak

Distilled Sensing: Adaptive Sampling for Sparse Detection and Estimation. 2010. AWGN, $Y_{i, j}=x_{i}+\gamma_{i, j}^{-1 / 2} W_{i, j}$, FDP/NDP
L. Lai, H. Vincent Poor, Y. Xin, and G. Georgiadis

Quickest Search Over Multiple Sequences. Trans. on Info Theory. 2011. Find one element in $\mathcal{S}$
E. Bashan, G. Newstadt, and A. Hero

Two-Stage Multi-Scale Search for Sparse Targets. 2011. AWGN, two stage procedure
A. Tajer, R. Castro

Adaptive Spectrum Sensing for Agile Cognitive Radios. 2010.
Spectrum Sensing


## Motivation - Gene knock-out studies in biology

fruit fly


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What can we do with limited knowledge of distributions?

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2. Sequential

- sequential thresholding or coordinate-wise SPRT [A. Tajer 2010, W. Zhang 2010, M. Malloy 2011]


## Limitations of non-sequential testing

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T_{i}^{(m)}:=\frac{1}{m} \sum_{j=1}^{m} \log \frac{p_{1}\left(Y_{i, j}\right)}{p_{0}\left(Y_{i, j}\right)}
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Non-sequential methods cannot overcome statistical 'curse of dimensionality'


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## The Sequential Probability Ratio Test

How much better can sequential methods do? [Wal, Wolfowit, Opimum Char. of the SPRT, 1948] On a coordinate-wise basis, take additional measurement of index $i$ if:

$$
A<\prod_{j=1}^{j^{\prime}} \frac{p_{1}\left(Y_{j}\right)}{p_{0}\left(Y_{j}\right)}<B
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$$
\log \frac{(n-s)}{\epsilon}
$$

$$
\log \frac{\epsilon}{s}
$$

Theorem
Set $A=\frac{\epsilon}{2 s}, B=\frac{2(n-s)}{\epsilon}$. The SPRT recovers $\mathcal{S}$ with probability

$$
\mathbb{P}(\hat{\mathcal{S}}=\mathcal{S}) \geq 1-\epsilon
$$

and requires fewer than

$$
m \leq\left(1+\epsilon_{0}\right) \frac{\log s+\log \epsilon^{-1}}{D\left(P_{0} \| P_{1}\right)}
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samples per dimension in expectation for any $\epsilon_{0}>0, n$ sufficiently large.


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Theorem
There exist thresholds $A$ and $B$ such that the SPRT recovers $\mathcal{S}$ exactly if

$$
m>\frac{\log s}{D\left(P_{0} \| P_{1}\right)}+\frac{\log \epsilon_{n}^{-1}}{D\left(P_{0} \| P_{1}\right)}
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for any $\epsilon_{n} \rightarrow 0$.


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- 1) for $s \ll n$
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m \approx \mathbb{E}_{0}[J]=\frac{\mathbb{E}_{0}\left[\sum_{j=1}^{J} \log \frac{p_{0}\left(y_{j}\right)}{p_{1}\left(y_{j}\right)}\right]}{D\left(P_{0} \| P_{1}\right)}
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- 2) Wald's identity
- 3) log-sum inequality for any test with false negative $\beta$, false positive $\alpha$
- 4) assume $\alpha \leq 1 /(n-s)$ and $\beta \leq 1 / s$
- 5) imposing condition

$$
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Sketch of coordinate-wise lower bound

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$12 / 17$

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- thresholds depend on $\mu$, number of samples
- What can we do without knowledge of $P_{1}$ or $s$ ?



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Input:

- $K \approx \log n$ measurement passes
- threshold $\gamma: \mathbb{P}_{0}\left(T^{(m / 2)} \leq \gamma\right)=\frac{1}{2}$
$13 / 17$


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1) sample each index $\frac{m}{2}$ times
2) re-measure only indices above threshold

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\mathbb{E}\left[\sum_{i=1}^{n} J_{i}\right] \approx \frac{m n}{2}+\frac{m n}{4}
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Sequential Thresholding succeeds in exactly recovery of $\mathcal{S}$ if

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Note: For certain levels of sparsity, ST is asymptotically optimal!

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$15 / 17$

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$16 / 17$

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## Conclusion

Remaining questions: can procedures remove doubly logarithmic gap without full knowledge of distributions?

For further reading:
围 M. Malloy, R. Nowak
Sequential Analysis in High Dimensional Multiple Testing and Sparse Recovery. ISIT 2011.
睩 M. Malloy, R. Nowak
On the limits of Sequential Testing in High Dimensions.
Asilomar 2011.


