

# Asymptotics for Causal Linear Fields

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Joint work with Neville Weber, University of Sydney



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A stationary process  $(X_i : i \in \mathbf{Z})$  or stationary random field  $(X_{i,j} : i, j \in \mathbf{Z})$  has short memory (is short-range dependent) if and only if its covariance function is absolutely summable:

$$\sum_i |\text{Cov}(X_0 X_i)| < \infty, \text{ resp.}, \sum_i \sum_j |\text{Cov}(X_{0,0} X_{i,j})| < \infty$$

Most known asymptotic results for short memory fields involve conditions on mixing or association. These are not needed in the case of a causal linear process (field).





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Let  $(\xi_i : i \in \mathbf{Z})$  be i.i.d. with mean 0 and variance 1 and let  $(a_i : i \in \mathbf{Z})$  be square summable:  $\sum_i a_i^2 < \infty$ .

$$X_s := \sum_i a_i \xi_{s-i}$$

is a linear process.

Let  $(\xi_{i,j} : i, j \in \mathbf{Z})$  be i.i.d. with mean 0 and variance 1 and let  $(a_{i,j} : i, j \in \mathbf{Z})$  be square summable:  $\sum_i \sum_j a_{i,j}^2 < \infty$ .

$$X_{s,t} := \sum_i \sum_j a_{i,j} \xi_{s-i,t-j}$$

is a linear field.





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- The linear process (field) has short memory if and only if  $\sum_i |a_i| < \infty$  (resp.,  $\sum_i \sum_j |a_{i,j}| < \infty$ ).
- The linear process (field) is *causal* if  $a_i = 0 \forall i < 0$  (resp.,  $a_{i,j} = 0$  if  $i < 0$  or  $j < 0$ ) i.e.

$$X_s := \sum_{i=0}^{\infty} a_i \xi_{s-i}$$

$$X_{s,t} := \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} a_{i,j} \xi_{s-i,t-j}$$

- Causal linear fields have received attention in the recent economics literature.





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Let  $F$  denote the marginal distribution of  $X_i$  (resp.,  $X_{i,j}$ ).  
We will consider the asymptotic behaviour of the  
corresponding empirical distributions:

$$F_n(x) := \frac{1}{n} \sum_{i=1}^n I(X_i \leq x)$$

$$F_{m,n}(x) := \frac{1}{mn} \sum_{i=1}^m \sum_{j=1}^n I(X_{i,j} \leq x)$$





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## Theorem

(*Doukhan and Surgailis (1998)*) Let  $(X_i)$  be a short memory linear process and assume that the  $\xi'_i$ 's have a bounded density. For each  $x \in \mathbf{R}$

$$\sqrt{n}(F_n(x) - F(x)) \xrightarrow{\mathcal{D}} W(x)$$

where  $W(x) \sim N(0, \sigma^2(x))$  and

$$\sigma^2(x) := \sum_{i=-\infty}^{\infty} \text{Cov}(I(X_0 \leq x), I(X_i \leq x)).$$







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If  $E[\xi_0^4] < \infty$ , then

$$\sqrt{n}(F_n(\cdot) - F(\cdot)) \xrightarrow{\mathcal{D}} W$$

in  $D(\mathbf{R})$  where  $W$  is a centered Gaussian process with covariance function

$$\sigma(x, y) := \sum_{i=-\infty}^{\infty} \text{Cov}(I(X_0 \leq x), I(X_i \leq y)).$$





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Let  $\mathcal{F}_i = \sigma\{\xi_u : u \leq i\}$ . By causality,  $X_i$  is  $\mathcal{F}_i$ -measurable.

$$U_i(h) := P[X_i \leq x | \mathcal{F}_{i-h}] - P[X_i \leq x | \mathcal{F}_{i-h-1}]$$

are martingale differences.

- $U_i(h)$  is  $\mathcal{F}_{i-h}$ -measurable.





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$$U_i(h) := P[X_i \leq x | \mathcal{F}_{i-h}] - P[X_i \leq x | \mathcal{F}_{i-h-1}]$$

are martingale differences.

- $U_i(h)$  is  $\mathcal{F}_{i-h}$ -measurable.
- For  $h$  fixed,  $U_i(h)$  is stationary in  $i$ .



# Gordin's Method (Gordin (1969))



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$$U_i(h) := P[X_i \leq x | \mathcal{F}_{i-h}] - P[X_i \leq x | \mathcal{F}_{i-h-1}]$$

are martingale differences.

- $U_i(h)$  is  $\mathcal{F}_{i-h}$ -measurable.
- For  $h$  fixed,  $U_i(h)$  is stationary in  $i$ .
- $\bigcap_i \mathcal{F}_i = \mathcal{T}$  satisfies the 0-1 law.



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For  $K < \infty$ ,

$$R_i^K(x) := \sum_{h=0}^K U_i(h) = I(X_i \leq x) - P[X_i \leq x | \mathcal{F}_{i-K-1}].$$

①  $R_i^K(x) \xrightarrow{a.s.} R_i(x) := I(X_i \leq x) - F(x)$  as  $K \rightarrow \infty$ .



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For  $K < \infty$ ,

$$R_i^K(x) := \sum_{h=0}^K U_i(h) = I(X_i \leq x) - P[X_i \leq x | \mathcal{F}_{i-K-1}].$$

- 1  $R_i^K(x) \xrightarrow{a.s.} R_i(x) := I(X_i \leq x) - F(x)$  as  $K \rightarrow \infty$ .
- 2  $\sqrt{n}(F_n(x) - F(x)) = \frac{1}{\sqrt{n}} \sum_{i=1}^n R_i(x) \approx \frac{1}{\sqrt{n}} \sum_{i=1}^n R_i^K(x)$ .



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$$\begin{aligned} \frac{1}{\sqrt{n}} \sum_{i=1}^n R_i^K(x) &= \frac{1}{\sqrt{n}} \sum_{i=1}^n \sum_{h=0}^K U_i(h) \\ &\approx \frac{1}{\sqrt{n}} \sum_{i=1}^n \sum_{h=0}^K U_{i+h}(h) \\ &= \frac{1}{\sqrt{n}} \sum_{i=1}^n M_i^K \end{aligned}$$

$\{M_i^K := \sum_{h=0}^K U_{i+h}(h) : i \geq 1\}$  is a *stationary and ergodic* sequence of martingale differences.



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$$\begin{aligned} E[(M_0^K)^2] &= \sum_{i=-\infty}^{\infty} \text{Cov}(R_0^K(x), R_i^K(x)) \\ &:= \sigma_K^2(x) \xrightarrow{K} \sigma^2(x). \end{aligned}$$

CLT for stationary ergodic martingale differences  $\Rightarrow$

$$\frac{1}{\sqrt{n}} \sum_{i=1}^n R_i^K(x) \xrightarrow{D} W_K(x)$$

where  $W_K(x) \sim N(0, \sigma_K^2(x))$ . Now let  $K \rightarrow \infty$ .

Fourth moments are used to prove tightness. □







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## Theorem

(Ivanoff and Weber (2010)) Let  $(X_{i,j})$  be a short memory linear field and assume that the  $\xi'_{i,j}$ s have a bounded density. For each  $x \in \mathbf{R}$

$$\sqrt{mn}(F_{m,n}(x) - F(x)) \xrightarrow{\mathcal{D}} W(x)$$

as  $m, n \rightarrow \infty$  where  $W(x) \sim N(0, \sigma^2(x))$  and

$$\sigma^2(x) := \sum_{i=-\infty}^{\infty} \sum_{j=-\infty}^{\infty} \text{Cov}(I(X_{0,0} \leq x), I(X_{i,j} \leq x)).$$





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If  $E[\zeta_{0,0}^4] < \infty$ , then as  $m, n \rightarrow \infty$

$$\sqrt{mn}(F_{m,n}(\cdot) - F(\cdot)) \xrightarrow{\mathcal{D}} W$$

in  $D(\mathbf{R})$  where  $W$  is a centered Gaussian process with  
covariance function  $\sigma(x, y) :=$

$$\sum_{i=-\infty}^{\infty} \sum_{j=-\infty}^{\infty} \text{Cov}(I(X_{0,0} \leq x), I(X_{i,j} \leq y)).$$





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The approach: find a *total* order  $\prec$  on the lattice that allows us to create a (one-dimensional) martingale difference sequence that has some sort of stationarity and ergodicity properties.

$\prec$  must be consistent with the usual *partial* order  $\leq$  on  $\mathbf{Z}^2$   
(  $i \leq i'$  and  $j \leq j'$   $\Leftrightarrow$   $(i, j) \leq (i', j')$   $\Rightarrow$   $(i, j) \preceq (i', j')$  )

Using  $\prec$ , we must be able to

- to count backward on  $(-\infty, i] \times (-\infty, j]$  to  $(-\infty, -\infty)$
- to count forward on  $\mathbf{Z}_+^2$  to  $(\infty, \infty)$





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Solution: We use the *diagonal* order:

$$(i, j) \prec (i', j') \Leftrightarrow i + j < i' + j'$$

or  $i + j = i' + j'$  and  $i < i'$ .

- $(i, j) \leq (i', j') \Rightarrow (i, j) \preceq (i', j')$  (consistency)
- In the diagonal order  $\preceq$ , for the immediate predecessor of  $(i, j)$  is  $(i - 1, j + 1)$ .
- However, for  $r \leq i$ ,

$$\sup\{(h, k) : (h, k) \prec (r, j), (h, k) \leq (i, j)\} = (i, r + j - i - 1)$$



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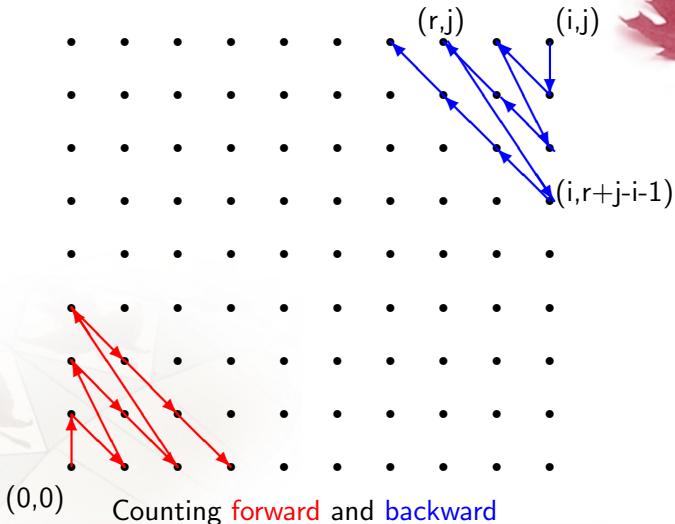
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Remember: the  $\xi'_{i,j}$ s are i.i.d.  
Define

$$\mathcal{F}_{i,j} = \sigma\{\xi_{u,v} : (u,v) \leq (i,j)\}$$

$$\mathcal{G}_{i,j} = \sigma\{\xi_{u,v} : (u,v) \preceq (i,j)\}$$

$$\mathcal{G}_\ell^D = \sigma\{\xi_{u,v} : u+v \leq \ell\}.$$

- By causality  $X_{i,j}$  is  $\mathcal{F}_{i,j}$ -measurable.
- Since  $\preceq$  is consistent with  $\leq$ ,  $\mathcal{F}_{i,j} \subseteq \mathcal{G}_{i,j}$ .





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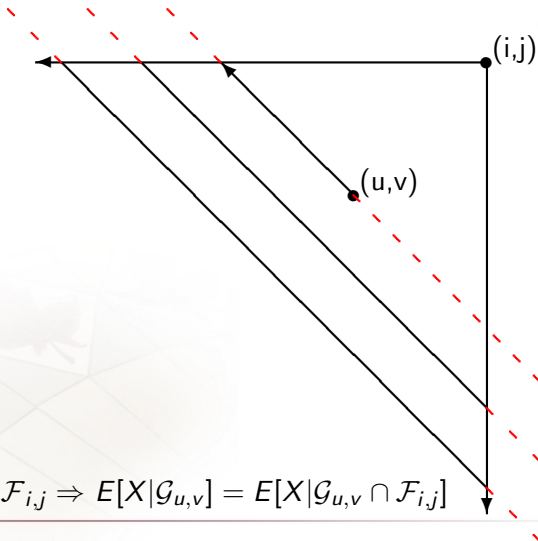
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$$X \in \mathcal{F}_{i,j} \Rightarrow E[X|\mathcal{G}_{u,v}] = E[X|\mathcal{G}_{u,v} \cap \mathcal{F}_{i,j}]$$





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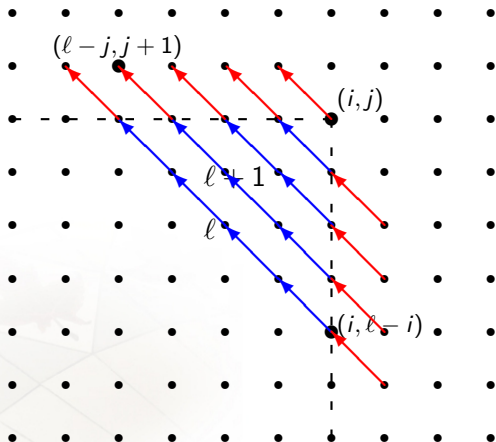
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$$E[X|\mathcal{G}_{l-j,j+1}] = E[X|\mathcal{G}_\ell^D] = E[X|\mathcal{G}_{i,l-i}]$$







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## Some observations:

If  $X$  is  $\mathcal{F}_{i,j}$  measurable then by independence of the  $\xi_{i,j}$ 's,

① if  $(u, v) \preceq (i, j)$

$$\begin{aligned} E[X|\mathcal{G}_{u,v}] &= E[X|\xi_{h,\ell}; (h, \ell) \preceq (u, v)] \\ &= E[X|\xi_{h,\ell}; (h, \ell) \preceq (u, v), (h, \ell) \leq (i, j)] \\ &= E[X|\mathcal{G}_{u,v} \cap \mathcal{F}_{i,j}]. \end{aligned}$$





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If  $X$  is  $\mathcal{F}_{i,j}$  measurable then by independence of the  $\xi_{i,j}$ 's,

① if  $(u, v) \preceq (i, j)$

$$\begin{aligned} E[X|\mathcal{G}_{u,v}] &= E[X|\xi_{h,\ell}; (h, \ell) \preceq (u, v)] \\ &= E[X|\xi_{h,\ell}; (h, \ell) \preceq (u, v), (h, \ell) \leq (i, j)] \\ &= E[X|\mathcal{G}_{u,v} \cap \mathcal{F}_{i,j}]. \end{aligned}$$

② if  $\ell < i + j$ , then

$$E[X|\mathcal{G}_{\ell-j, j+1}] = E[X|\mathcal{G}_{\ell}^D] = E[X|\mathcal{G}_{\ell}^D \cap \mathcal{F}_{i,j}] = E[X|\mathcal{G}_{i, \ell-i}]$$



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Recall:

$$X_{i,j} := \sum_{u=0}^{\infty} \sum_{v=0}^{\infty} a_{u,v} \xi_{i-u,j-v}$$

For all  $(i, j) \in \mathbf{Z}^2$  and  $h, k \geq 0$  define

$$U_{i,j}(h, k) := P[X_{i,j} \leq x | \mathcal{G}_{i-h,j-k}] - P[X_{i,j} \leq x | \mathcal{G}_{i-h-1,j-k+1}],$$

noting that if  $k = 0$ ,

$$\begin{aligned} U_{i,j}(h, 0) &= P[X_{i,j} \leq x | \mathcal{G}_{i-h,j}] - P[X_{i,j} \leq x | \mathcal{G}_{i-h-1,j+1}] \\ &= P[X_{i,j} \leq x | \mathcal{G}_{i-h,j}] - P[X_{i,j} \leq x | \mathcal{G}_{i,j-h-1}] \end{aligned}$$



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- For  $h, k$  fixed, the  $U_{i,j}(h, k)$  are stationary in  $i, j$ .



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- For  $h, k$  fixed, the  $U_{i,j}(h, k)$  are stationary in  $i, j$ .
- $U_{i,j}(h, k)$  is  $\mathcal{G}_{i-h, j-k}$ -measurable.



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- For  $h, k$  fixed, the  $U_{i,j}(h, k)$  are stationary in  $i, j$ .
- $U_{i,j}(h, k)$  is  $\mathcal{G}_{i-h, j-k}$ -measurable.
- $\cap \mathcal{G}_\ell^D = \mathcal{T}$  satisfies the 0-1 law .



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Recall

$$F_{m,n}(x) := \frac{1}{mn} \sum_{i=1}^m \sum_{j=1}^n I(X_{i,j} \leq x)$$

Condition backwards in  $\prec$  from  $(i, j)$  over  $K$  diagonals:

$$\begin{aligned} R_{i,j}^K(x) &:= \sum_{\ell=0}^K \sum_{h=0}^{\ell} U_{i,j}(h, \ell - h) \\ &= I(X_{i,j} \leq x) - P(X_{i,j} \leq x | \mathcal{G}_{i+j-(K+1)}^D) \end{aligned}$$



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- $R_{i,j}^K(x) \xrightarrow{a.s} R_{i,j}(x) := I(X_i \leq x) - F(x)$  as  $K \rightarrow \infty$ .





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- $R_{i,j}^K(x) \xrightarrow{a.s.} R_{i,j}(x) := I(X_i \leq x) - F(x)$  as  $K \rightarrow \infty$ .
- 

$$\begin{aligned}\sqrt{mn}(F_{m,n}(x) - F(x)) &= \frac{1}{\sqrt{mn}} \sum_{i=1}^m \sum_{j=1}^n R_{i,j}(x) \\ &\approx \frac{1}{\sqrt{mn}} \sum_{i=1}^m \sum_{j=1}^n R_{i,j}^K(x)\end{aligned}$$



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$$\frac{1}{\sqrt{mn}} \sum_{i=1}^m \sum_{j=1}^n R_{i,j}^K(x)$$

$$= \frac{1}{\sqrt{mn}} \sum_{i=1}^m \sum_{j=1}^n \sum_{\ell=0}^K \sum_{h=0}^{\ell} U_{i,j}(h, \ell - h)$$

$$\approx \frac{1}{\sqrt{mn}} \sum_{i=1}^m \sum_{j=1}^n \sum_{\ell=0}^K \sum_{h=0}^{\ell} U_{i+h,j+\ell-h}(h, \ell - h)$$

$$= \frac{1}{\sqrt{mn}} \sum_{i=1}^m \sum_{j=1}^n M_{i,j}^K$$

$$\text{where } M_{i,j}^K = \sum_{\ell=0}^K \sum_{h=0}^{\ell} U_{i+h,j+\ell-h}(h, \ell - h).$$



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$\{M_{ij}^K : 1 \leq i \leq m, 1 \leq j \leq n\}$  is

- a stationary ergodic array (in two dimensions)



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$\{M_{ij}^K : 1 \leq i \leq m, 1 \leq j \leq n\}$  is

- a stationary ergodic array (in two dimensions)
- a sequence of identically distributed (but not stationary)  $mn$  martingale differences in the total order  $\preceq$  (in one dimension).



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A martingale CLT due to McLeish (1974) can now be applied.

- Since  $|U_{i,j}(h, k)| \leq 1$  and  $M_{i,j}^K$  contains  $\frac{(K+1)(K+2)}{2}$  summands,

$$\max_{1 \leq i \leq m, 1 \leq j \leq n} \frac{1}{\sqrt{mn}} |M_{i,j}^K|$$

is uniformly bounded and converges to 0 a.s.

- The two dimensional ergodic theorem implies that

$$\frac{1}{mn} \sum_{i=1}^m \sum_{j=1}^n (M_{i,j}^K)^2 \xrightarrow{P} E((M_{0,0}^K)^2).$$



# Gordin's Method in 2 Dimensions

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$$\begin{aligned} E((M_{0,0}^K)^2) &= \sum_{i \in \mathbf{Z}} \sum_{j \in \mathbf{Z}} \text{Cov}(R_{0,0}^K(x) R_{i,j}^K(x)) \\ &= \sigma_K^2(x) \rightarrow \sigma^2(x) \text{ as } K \rightarrow \infty \end{aligned}$$

The pointwise (and fidi) CLT follows assuming only second moments.

Fourth moments are used to prove tightness. □





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Gordin's martingale method can be extended to two dimensions via the diagonal order if

- the underlying filtration is generated by i.i.d. random variables, and
- the stationary process is causal

The technique combines one-dimensional martingale methods with two-dimensional ergodicity.





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