Thoughts on Model Selection

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Contributions

Thanks to

Lucas Kania,

Mikael Kuusela,

Nick Wardle, and

Larry Wasserman,

for their contributions to this talk.

Overview

"Statistician's view on model selection"

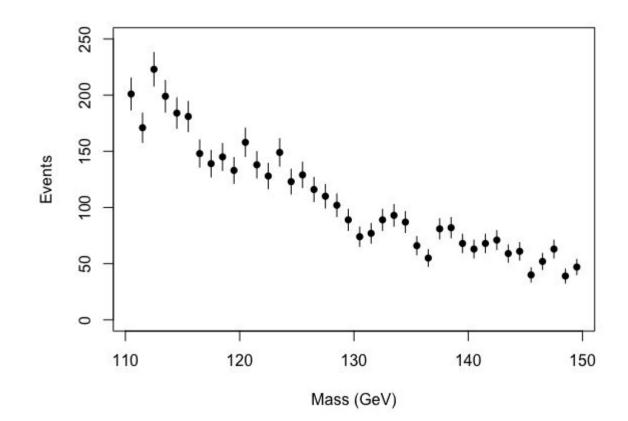
Very large topic

Touch on a few, relevant (I hope), concepts and ideas

- parametric versus nonparametric models
- AIC
- discrete profiling
- semiparametric models
- method of sieves
- model averaging

Focus on the specific problem of **background model selection** and the implications for **post-model selection inference**

Overview



Simulated data, courtesy of Nick Wardle

General form:

$$\lambda_i = \eta(m_i) + \theta N(125, 1.19)$$
$$X_i \sim \text{Poisson}(\lambda_i)$$

Consider polynomial background models, estimate via maximum likelihood

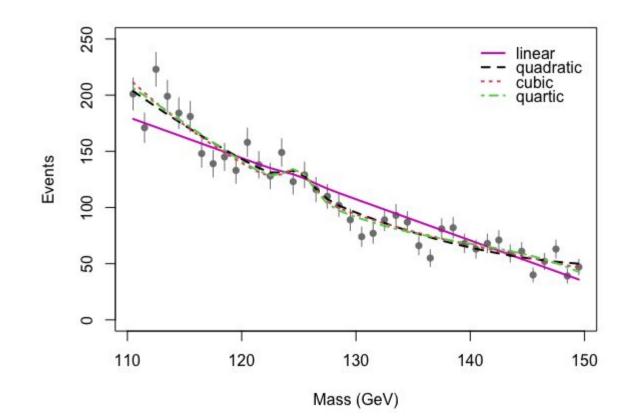
General form:

Background Signal

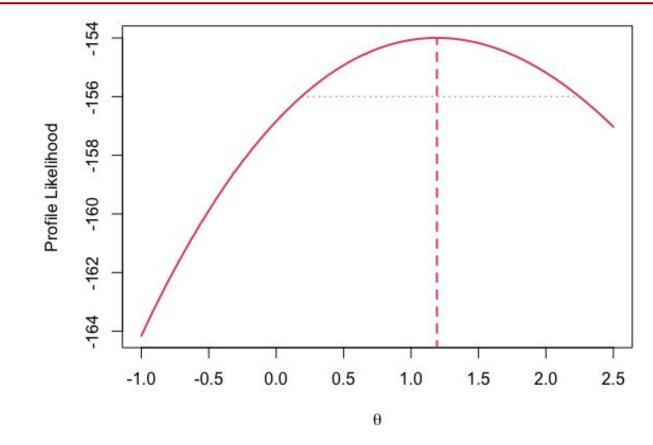
$$\lambda_i = \eta(m_i) + \theta N(125, 1.19)$$

 $X_i \sim \text{Poisson}(\lambda_i)$

Consider polynomial background models, estimate via maximum likelihood



Simulated data, courtesy of Nick Wardle



Profile likelihood for the quartic case

Model	MLE of θ	95% Cl for θ	Log likelihood at MLE	Number of Parameters	AIC
Linear	0.13	(-0.75, 1.12)	-174.79	3	355.58
Quadratic	0.84	(-0.14, 1.80)	-155.85	4	319.70
Cubic	1.14	(0.2, 2.10)	-154.81	5	319.62
Quartic	1.21	(0.23, 2.17)	-154.02	6	320.05

 $AIC = -2 \times log likelihood at max + 2 \times number of parameters$

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Quartic	1.21	(0.23, 2.17)	-154.02	6	320.05
Quintic	1.18	(0.29, 2.18)	-153.84	7	321.67
Sextic	1.24	(0.33, 2.19)	-153.83	8	323.65
Septic	1.25	(0.39, 2,30)	-153.34	9	324.69

Parametric versus Nonparametric

Is this a **parametric** or a **nonparametric** approach?

The process described above for choosing the background model is nonparametric.

The order of the model is serving as a **smoothing parameter**.

AIC is a widely used approach to choosing its value.

More data = less "smoothing"

Defines a de facto s**emiparametric approach**, where complexity of model is limited by ensuring consistency of estimation of θ

Semiparametric inference:

Parameter of interest: θ, lies in a Euclidean space

Nuisance parameter: η, lies in a more general space, denote it **N**

Murphy and van der Vaart (2000), "On Profile Likelihood":

"We show that semiparametric profile likelihoods, where the **nuisance parameter has been profiled out**, behave like ordinary likelihoods in that they have a quadratic expansion. In this expansion the score function and the Fisher information are replaced by the **efficient score function** and **efficient Fisher information**.

Murphy and van der Vaart (2000), "On Profile Likelihood":

"We show that semiparametric profile likelihoods, where the nuisance parameter has been profiled out, behave like ordinary likelihoods in that they have a quadratic expansion. In this expansion the score function and the Fisher information are replaced by the efficient score function and efficient Fisher information. The expansion may be used, among others, to prove the asymptotic normality of the maximum likelihood estimator, to derive the asymptotic chi-squared distribution of the log-likelihood ratio statistic, and to prove the consistency of the observed information as an estimator of the inverse of the asymptotic variance."

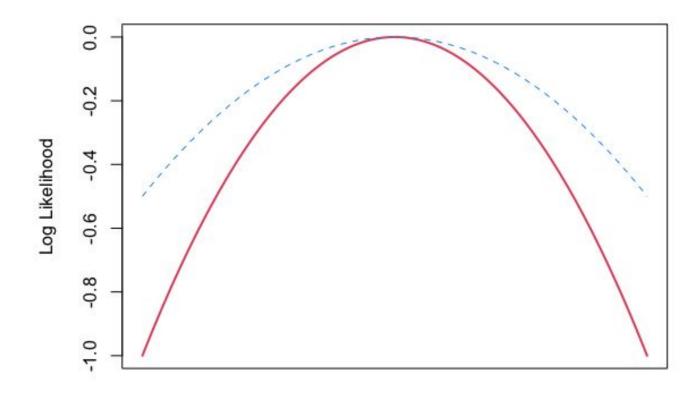
The score function for θ

$$l_{\theta_0,\eta_0} = \left. \frac{\partial}{\partial \theta} \log L(\theta,\eta) \right|_{\theta_0,\eta_0}$$

The efficient score function for θ

$$\tilde{l}_{\theta_0,\eta_0} = l_{\theta_0,\eta_0} - \Pi_{\theta_0,\eta_0} l_{\theta_0,\eta_0}$$

where Π projects onto the space of score functions for $\eta,$ finding the **least favorable model in N**

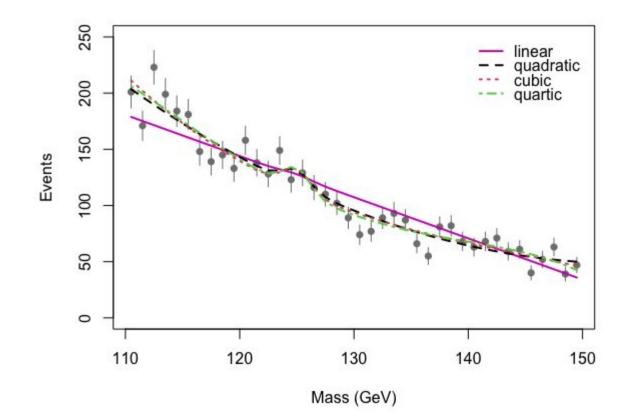


The efficient Fisher Information

$$\operatorname{Var}(\tilde{l}_{\theta_0,\eta_0}) = \tilde{\mathcal{I}}_{\theta_0,\eta_0}$$

The MLE for θ is approximately normal with covariance matrix

$$ilde{\mathcal{I}}_{\widehat{ heta},\widehat{\eta}}^{-1}$$



Simulated data, courtesy of Nick Wardle

Important Question:

What controls the complexity of the background? I.e., what limits N?

Could utilize physical constraints

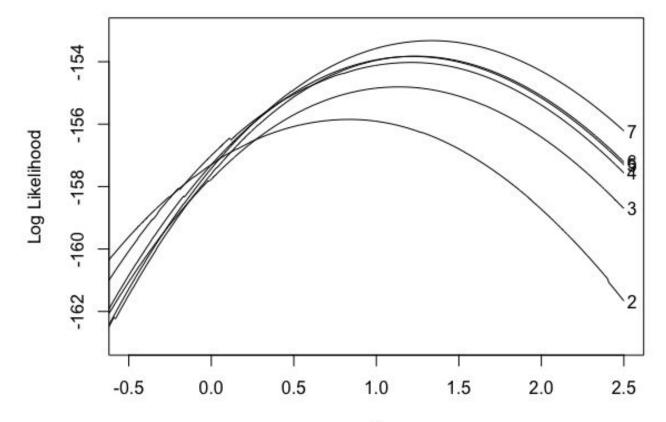
Possibly use Method of Sieves

(Grenander 1981, Geman and Hwang 1982)

Construct **series** of N_n which grow in complexity with n, but slowly enough to ensure **consistency** of estimating θ

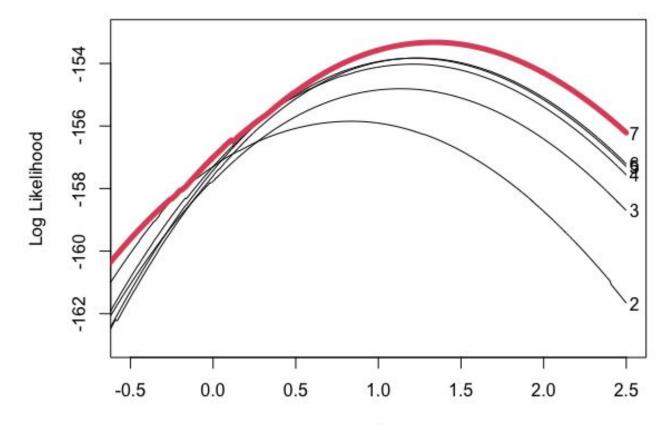
Interesting parallel with **Discrete Profiling** (Dauncey, et al. 2015)

Discrete Profiling



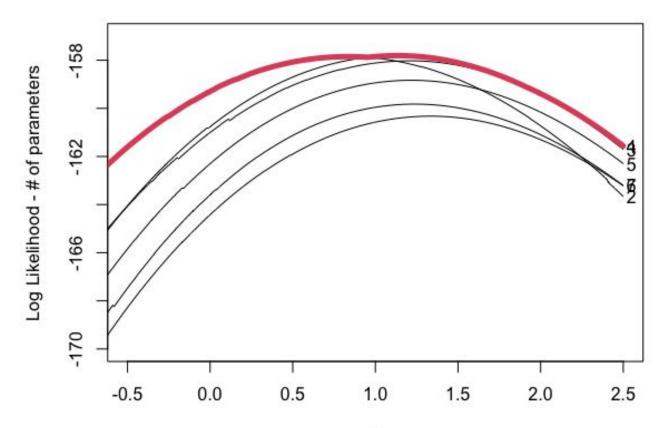
θ

Discrete Profiling



θ

Discrete Profiling



θ

Another class of ideas: Model Averaging

Overview in Chapter 7 of Claeskens and Hjort (2008)

Also, Burnham and Anderson (2002)

Instead of fixing on one model, average over multiple candidates

Buckland, et al. (1997):

General form for **Information criterion**:

$$I_k = -2\log(L_k) + q_k$$

where q_k is a penalty term.

Then, define model weights:

$$w_i \propto \exp(-I_i/2) = L_i \exp(-q_i/2)$$

General form for Information criterion:

$$I_k = -2\log(L_k) + q_k$$

where q_k is a penalty term.

For example, if

$$q_k = \log(n)p_k$$

then using **BIC**, and (see Schwarz 1978):

 $\frac{w_i}{w_j} \approx \text{Bayes factor between models } i \text{ and } j$

General form for Information criterion:

$$I_k = -2\log(L_k) + q_k$$

where q_k is a penalty term.

For example, if

$$q_k = 2p_k$$

then using AIC, and construct Akaike weights.

Model	MLE of θ	Log likelihood at MLE	AIC	Akaike Weight
Linear	0.13	-174.79	355.58	≈ 0
Quadratic	0.84	-155.85	319.70	0.29
Cubic	1.14	-154.81	319.62	0.30
Quartic	1.21	-154.02	320.05	0.24
Quintic	1.18	-153.84	321.67	0.11
Sextic	1.24	-153.83	323.65	0.040
Septic	1.25	-153.34	324.69	0.024

The **weighted estimator** for the parameter:

$$\widehat{\theta}_a = \sum_{i=1}^K w_i \widehat{\theta}_i$$

with variance (Equation 4.9 in Burnham and Anderson):

$$\operatorname{Var}\left(\widehat{\theta}_{a}\right) = \left[\sum_{i=1}^{K} w_{i} \sqrt{\operatorname{Var}\left(\widehat{\theta}_{i} | \text{model i}\right) + \left(\widehat{\theta}_{i} - \widehat{\theta}_{a}\right)^{2}}\right]^{2}$$

Model	MLE of θ	Log likelihood at MLE	AIC	Akaike Weight
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Quartic	1.21	-154.02	320.05	0.24
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 $1.08 \pm 1.96 \times 0.050 = (0.09, 2.07)$

Improved versions of confidence intervals in

Section 4.3.3 in Burnham and Anderson

Chapter 7 of Claeskens and Hjort

References

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